

ORIGINAL ARTICLE

Bridging disconnected networks of first and second lines of biologic therapies in rheumatoid arthritis with registry data: bayesian evidence synthesis with target trial emulation

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Abstract

Objectives: We aim to use real-world data in evidence synthesis to optimize an evidence base for the effectiveness of biologic therapies in rheumatoid arthritis to allow for evidence on first-line therapies to inform second-line effectiveness estimates.

Study Design and Setting: We use data from the British Society for Rheumatology Biologics Register for Rheumatoid Arthritis to supplement randomized controlled trials evidence obtained from the literature, by emulating target trials of treatment sequences to estimate treatment effects in each line of therapy. Treatment effects estimates from the target trials inform a bivariate network meta-analysis (NMA) of first-line and second-line treatments.

Results: Summary data were obtained from 21 trials of biologic therapies including two for second-line treatment and results from six emulated target trials of both treatment lines. Bivariate NMA resulted in a decrease in uncertainty around the effectiveness estimates of the second-line therapies, when compared to the results of univariate NMA, and allowed for predictions of treatment effects not evaluated in second-line randomized controlled trials.

Conclusion: Bivariate NMA provides effectiveness estimates for all treatments in first and second line, including predicted effects in second line where these estimates did not exist in the data. This novel methodology may have further applications; for example, for bridging networks of trials in children and adults. © 2022 The Author(s). Published by Elsevier Inc. This is an open access article under the CC BY license (<http://creativecommons.org/licenses/by/4.0/>).

Keywords: Real world evidence; Bivariate network meta-analysis; Target trial emulation; Treatment lines; Rheumatoid arthritis; Biologic therapies

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What is new?**Key findings**

- Using real world evidence (RWE) on the effectiveness of biologic therapies in rheumatoid arthritis, by applying a target trial emulation approach and Bayesian evidence synthesis, we allowed for evidence on first-line therapies to inform second-line effectiveness estimates.

What this adds to what was known?

- Traditionally, data from randomised controlled trials (RCTs) have been used to inform evidence based decision making; however RCTs are typically carried out in either first or second line of therapy.
- RWE increasingly is considered a valuable source of evidence and it may offer long term follow up of patients and their outcomes from multiple lines of treatments.

What is the implication and what should change now?

- Using RWD to make inferences about different lines of therapy may aid trial design and potentially policy decisions.

second-line effectiveness estimates in evidence synthesis. When data from RCTs are available on effectiveness of a particular treatment, but only in the first line of therapy, a costly trial needs to be undertaken to also evaluate the effectiveness of the new therapy used in patients as a second-line treatment (or vice versa). We investigated the added value of registry data, which provides evidence on both first and second lines in each individual, when amalgamating these data in a network of RCTs for both lines of therapies. We developed this approach for incorporating RWD into clinical and health technology assessment decision-making using a case study in rheumatoid arthritis (RA).

We made use of data from the British Society for Rheumatology Biologics Register for Rheumatoid Arthritis (BSRBR-RA) to supplement the RCT evidence available only for either first or the second line of therapy. We did so by emulating target trials using the approach developed by Hernán and Robins [7]. We estimated treatment effects of biologic therapies based on the data in emulated target trials, which we then used to inform a bivariate network meta-analysis (NMA) model of first-line and second-line treatments. The estimates from the registry data were used to “bridge” disconnected networks for the two lines of therapy. The American College of Rheumatology response criteria (ACR20) were used as an outcome measure.

The remainder of this article is structured as follows. Data sources and statistical methods are described in Section 2. The results are presented in Section 3, which are followed by discussion and conclusion in Section 4.

1. Introduction

The evidence base for healthcare decision-making traditionally consisted of data from randomized controlled trials (RCTs), considered as a gold standard in evaluation of health technologies. In recent years, there has been a growing interest in the use of real-world data (RWD) from observational studies in healthcare evaluation. Routinely collected data, from electronic health records or patients’ registries, can provide useful information about effectiveness of treatments, where data from RCTs may be sparse or are not available at all for some treatment comparisons. Considerable methodological research has focused on inclusion of RWD in evidence synthesis with the aim of overcoming some limitations of RCT data [1–3]. The focus of such research has been particularly in circumstances where RCT evidence was sparse and combining RCT data with RWD aimed to increase the evidence base to improve the precision of effectiveness estimates [4] and sometimes bridge disconnected networks.

While research to date has largely focused on exploitation of RWD to mimic or replicate RCT data [5–7], we take a step further to explore the use of RWD in a scenario of data generation not typical for the RCT setting. In this article, we explored how RWD can be used to optimize an evidence base by using evidence on first-line therapies to inform

2. Methods*2.1. Data sources**2.1.1. Summary data from randomized controlled trials*

Summary data from a literature review of RCTs of biologic therapies in patients with RA were obtained for the effectiveness of adalimumab, etanercept, infliximab, golimumab, abatacept, and rituximab used as first-line biologic therapies (in biologic naive patients) and the effectiveness of golimumab and rituximab used as second-line biologic therapies in patients who switched from a previous biologic treatment. Data were obtained from 20 trials including 18 for the first-line treatments and two trials for second-line treatments. A list of references for the trials is included in the [Supplementary Materials A](#). When constructing a network, placebo arms with methotrexate as concomitant therapy and the arms including a combination of methotrexate and placebo were treated as the same treatment arm in the network. Methotrexate, used in many trials as part of the combination therapy in the biologic arm, was ignored (for some studies methotrexate was included as concomitant therapy where percentage of patients with addition of methotrexate varied across studies, similarly as in the BSRBR-RA target trials).

2.1.2. Registry data

We made use of data from the BSRBR-RA to supplement randomized trial evidence. While RCTs included only either first-line or second-line therapy, registry data provided evidence on both lines of therapy for each patient. BSRBR-RA data consisted of 19,410 individuals, 15,636 of whom had data recorded on biologic treatment. The data were used to emulate target trials of both lines of therapy.

2.2. Emulation of target trials

We used the BSRBR-RA data to emulate a series of trials of first-line and second-line treatments for a range of biologic therapies using a target trial approach [7]. In the first instance, we specified the key components of the target trial protocol, which (following recommendation by Hernán and Robins [7]) included eligibility criteria, treatment strategies, assignment procedures, the follow-up period, outcome, causal contrast, and statistical analysis. Note that no RCT included both lines of therapy in sequence, while the proposed protocol of the target trial did include the treatment sequence. Therefore, we did not aim for the emulated target trials to resemble any existing RCT (an approach previously used in target trial emulation).

2.2.1. Eligibility criteria

Study participants were aged 18 years or older, who had a diagnosis of RA. Patients who were treated with a biologic disease-modifying antirheumatic drug (DMARD) prior to the registration with the BSRBR-RA were excluded.

2.2.2. Treatment strategies

Patients had to have received at least two lines of therapy, which could be any of the biologic DMARDs or methotrexate, which is a synthetic DMARD often used as a combination therapy and/or control treatment in trials of biologic therapies in RA patients. Data from patients who switched from first-line biologic therapy to no therapy (or to therapies that are neither biologic DMARDs nor methotrexate) were not included.

2.2.3. Assignment procedures

Patients were grouped into treatment arms as per the sequence of treatment in two lines of therapy. These groups of patients (sequence treatment arms) were matched to form experimental and control treatment groups. Matching was conducted based on size of the trial, ensuring well-balanced treatment contrasts, with methotrexate always taken as the control treatment and rituximab as an experimental treatment. Other biologic therapies could be used as either experimental treatment or control. The matching procedure had to ensure unique treatments in experimental and control arms for each line. The process is schematically described in [Figure 1](#). This procedure resulted in target trials of two lines of therapy recorded on the same patients who switched treatment in both treatment arms. For

example, patients in the experimental arm receiving first-line adalimumab switched to infliximab and those in the control arm receiving first-line etanercept switched to methotrexate, thus resulting in the first-line comparison of adalimumab vs. etanercept and in the second-line comparison of infliximab vs. methotrexate. Since patients were not randomly allocated, we assumed no unmeasured confounding at baseline conditional on a number of prognostic factors measured at baseline or initiation of each treatment that could influence the response. The prognostic factors included age, gender, duration of the disease, serology (being positive for rheumatoid factor), and 28-joint count disease activity score (DAS-28).

2.2.4. The follow-up period

The minimum follow-up time had to ensure that data were collected 24 weeks after initiation of each line of therapy. Start of the second-line therapy varied depending when patients needed to switch to second-line treatment, which was typically due to either a lack of response or adverse reactions.

2.2.5. Outcome

Patients were assessed as per ACR20 response criteria, which classified them as responders if they had at least 20% improvement as per ACR criteria. Due to a large number of missing values on some of the components of ACR within BSRBR-RA data (the register did not capture patient pain or physician global score), the definition of response was relaxed allowing patients to be classed as responders if they had at least 20% improvement in at least one of the joint count components (tender or swollen joint count) and at least one of the remaining five components of the ACR measure (physician global assessment, patient global assessment, pain, health assessment questionnaire, erythrocyte sedimentation rate [or C-reactive protein]) [8].

2.2.6. Causal contrast and statistical analysis

Baseline characteristics for each group were summarized to ensure that the covariates were similarly distributed across the treatment arms. The numbers of responders were then adjusted for covariates using inverse probability weighting with propensity scoring (IPW-PS) [9,10]. A schematic causal diagram with potential confounders, taken into account in the analysis, is included in [Figure B1](#) of the [Supplementary materials B](#). We estimated the per-protocol effect in all emulated trials.

2.3. Bivariate network meta-analysis

A univariate random effect NMA was used to model data on second line of therapy using RCT data alone, BSRBR-RA data alone (with both adjusted and unadjusted BSRBR-RA data), and combined data from RCTs and the register. We then used bivariate NMA to model jointly the treatment effects on ACR20 for first and second lines

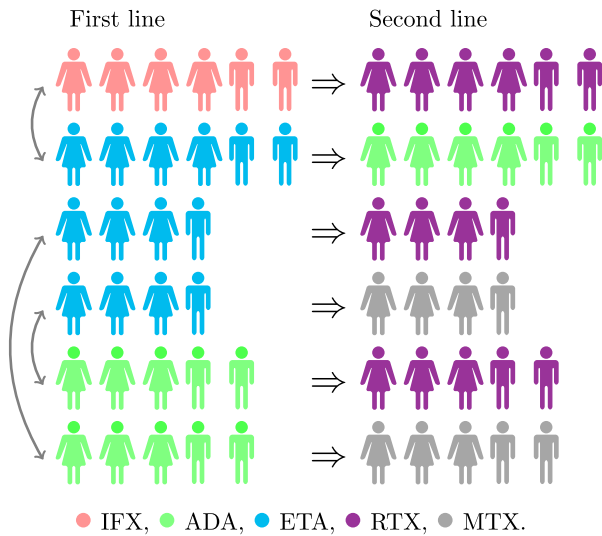


Fig. 1. Schematic diagram representing the process of matching sequence treatment arms. Each row represents patients assigned to a unique treatment sequence depicted by different colors. Gray arrows on the left hand side show how the sequence treatment groups were matched to ensure a balance in terms of the size of each arm and that for both lines the treatment in each arm was different. ADA, adalimumab; ETA, etanercept; IFX, infliximab; RTX, rituximab; MTX, methotrexate.

of therapies. A standard approach to any multivariate meta-analysis is to use a hierarchical model with a multivariate normal distribution used to describe variability at two levels: within-study (where the correlation occurs due to the modeled multivariate quantities, such as treatment effects on multiple outcomes, being measured in the same individuals) and between-studies (where the correlation is a result of heterogeneity between the average effects, measured on each outcome in each study, varying across studies due to, for example, differences in population or treatment doses). Accounting for the within-study correlation is important in such analysis [11]. However, modeling jointly non-normal outcomes, such as binomial responses, would require transforming data, which can lead to biased results [12]. Papanikos et al. carried out a simulation study showing that when the within-study correlation is weak, a multivariate meta-analysis model with independent binomial likelihoods is preferable [13]. An exploratory analysis of the BSRBR-RA dataset, estimating the within-study correlation using the bootstrapping approach [14,15], showed that the within-study correlation between the treatment effects for the two lines of therapy transformed onto the log odds ratio (OR) scale was close to zero. We, therefore, adapted the approaches to multivariate/bivariate NMA by Achana et al. [16] and Bujkiewicz et al. [17] by assuming independent binomial likelihoods at the within-study level, as in Papanikos et al. [13], to model the proportions of responders to treatment in each line of therapy. To predict treatment effects in the second line when data are only available for the therapy in first line, additional assumptions of exchangeability needed to be made, where instead of placing prior distributions on basic parameters, we added

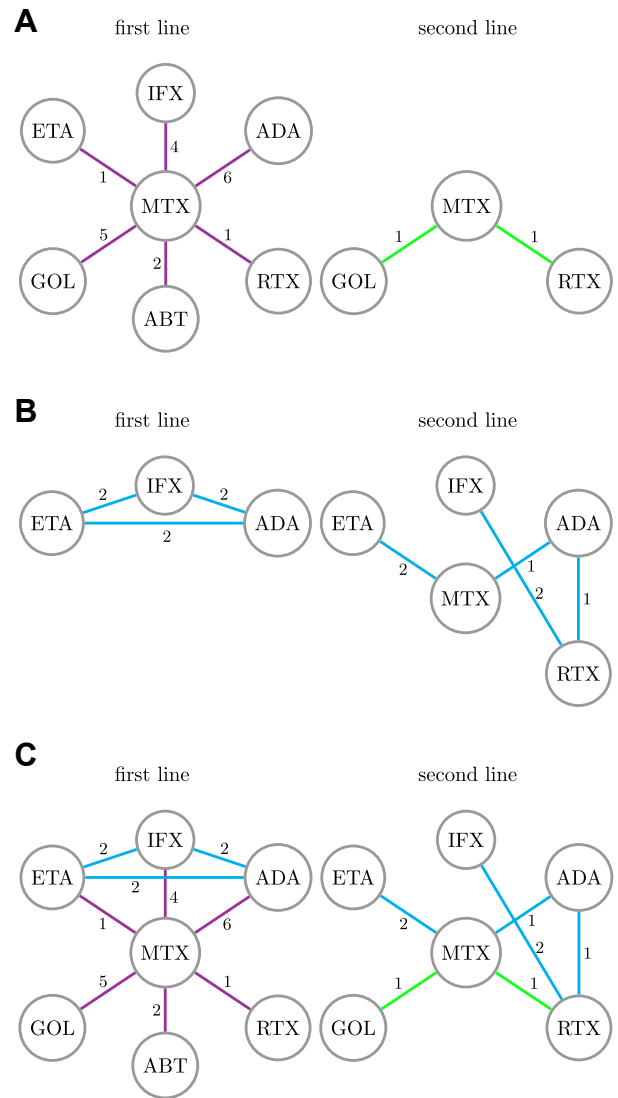


Fig. 2. Network diagram for (A) the RCT data, (B) BSRBR-RA data, and (C) combined data; first-line treatments (left) and second-line treatments (right).

another level of hierarchy to the model as in Bujkiewicz et al [17]. The details of both models are included in the [Supplementary materials C](#)

3. Results

3.1. Summary of data and the network structure

Summary data were obtained from 20 RCTs of biologic therapies with 18 trials for first-line treatment (including adalimumab, etanercept, infliximab, golimumab, abatacept, and rituximab) and two for second-line treatment (including golimumab and rituximab). BSRBR-RA data included 12,657 individuals given first-line biologic therapies at the time of registration. Follow-up data included 112,983 observations, which was on average 8.93 follow-ups per individual. For a large proportion of the visits,

Table 1. Results of a univariate fixed effects NMA of biologic therapies used as second-line treatments using data from RCTs alone, represented as ORs with 95% CrIs. Left-hand-side estimates correspond to the results of fixed effects NMA (upper triangle) and direct evidence (lower triangle) and the right-hand-side results are from the random effects NMA. ORs from the direct evidence are accompanied with 95% confidence intervals

Treatment	Fixed effects			Random effects		
	MTX	GOL	RTX	MTX	GOL	RTX
MTX		2.73 (1.68, 4.31)	4.84 (3.14, 7.25)	6.41 (0.21, 33.4)		11.34 (0.37, 59.4)
GOL	2.62 (1.64, 4.19)		1.88 (0.95, 3.36)			15.7 (0.05, 64.3)
RTX	4.68 (3.08, 7.1)	–				

methotrexate was recorded as a concomitant therapy to a biologic treatment. Target trial emulation using the BSRBR-RA data led to generation of six target trials of biologic therapies in two lines. The distribution of the covariates in the target trials are listed in Table B1 of [Supplementary materials B](#).

Figure 2a shows the network structure of RCT data for the first and second lines of therapy and Figure 2b illustrates the network structure of target trials emulated from BSRBR-RA data for both lines of therapy. For the target trials, both treatment lines correspond to the same trial, in contrast to the RCTs which report only either first or second line of treatment. To emphasize this in Figure 2, we used the same color of the network edges for both treatment lines for the target trials, in contrast to the RCTs where different colors of edges for different treatment lines represent different trials. In this article, we aimed to demonstrate the value of the registry data in estimating the effect of the biologic therapies when used as second-line treatments. The network of RCT data for the second-line therapy was particularly sparse, including only two trials, for golimumab and rituximab. BSRBR-RA data gave additional information about adalimumab, etanercept, infliximab, and rituximab used as second-line therapies. The network structure for RCT and BSRBR-RA data combined is shown in Figure 2c. The information on the number of studies for each treatment contrast, line of therapy, and the type of study is also included in Table B2 of the [Supplementary materials B](#).

3.2. Results of network meta-analyses

The results of all NMAs represent ORs with 95% credible intervals (CrIs), corresponding to ACR20 response. Results of NMAs of RCT data for the biologic therapies used

as second-line treatments are shown in Table 1. To comply with the GRADE guidelines [18], we compare the results from the NMA in the upper triangle (left-hand-side part of the table) with the direct effects in the lower triangle. The NMA results were obtained using the fixed-effects NMA model, as only two RCTs were included. On the right-hand-side of the table included are the results of the random effects meta-analysis for the purpose of like-with-like comparisons with the results of more complex analyses discussed below that used the random effects NMAs. Before combining the RCT data with the BSRBR-RA data, we carried out NMA of BSRBR-RA data alone, using both adjusted (with IPW-PS) and unadjusted data for comparison. The results of these analyses, shown in Table 2, did not appear very different, suggesting a good balance of covariates across the treatment arms in the target trials.

Results of two univariate NMAs of biologic therapies used as second-line treatments using combined data from the RCTs and BSRBR-RA (both adjusted and unadjusted) are shown in Table 3. Including the registry data allowed for estimation of treatment effects for second-line biologic therapies which were not included in the RCT network.

There was also some improvement in the precision of treatment effect estimates for those already included in the RCT data (when comparing the results from random effects NMA of the RCT data). For example, comparing rituximab with methotrexate gave OR = 11.3 (95% CrI: 0.4, 59.4) when using RCT data alone, while addition of the registry data resulted in OR = 3.9 (0.5, 14.9), thus reducing the uncertainty by 76% in terms of the width of the credible interval. This is expected considering inflated intervals around effects obtained from the random effects NMA of only two RCTs. The results of NMAs combining data from the RCTs and BSRBR-RA were, however, inflated when compared with those from NMA of the registry data alone.

Table 2. Results of a univariate NMA of biologic therapies used as second-line treatments using data from BSRBR-RA alone, adjusted using IPW-PS (upper triangle) and unadjusted (lower triangle), represented as ORs with 95% CrIs

Treatment	MTX	ADA	ETA	IFX	RTX
MTX		3.04 (0.31, 11.7)	3.94 (0.89, 11.7)	10.37 (0.08, 25.5)	2.87 (0.06, 10.5)
ADA	2.98 (0.32, 11.01)		3.77 (0.18, 15.5)	1.73 (0.08, 7.08)	0.68 (0.07, 2.64)
ETA	3.89 (0.91, 11.3)	3.74 (1.19, 14.9)		5.23 (0.02, 10.5)	1.45 (0.01, 4.44)
IFX	9.31 (0.09, 25.4)	1.8 (0.09, 7.28)	5.46 (0.02, 10.4)		0.7 (0.15, 2.06)
RTX	3.03 (0.07, 10.0)	0.69 (0.07, 2.56)	1.63 (0.02, 4.22)	0.66 (0.15, 1.92)	

Table 3. Results of a univariate NMA of biologic therapies used as second-line treatments using combined data from RCTs and BSRBR-RA adjusted using IPW-PS (upper triangle) and unadjusted (lower triangle), represented as ORs with 95% CrIs

Treatment	MTX	ADA	ETA	IFX	GOL	RTX
MTX		5.03 (0.59, 19.4)	4.26 (0.71, 14.4)	10.3 (0.45, 47.1)	5.1 (0.3, 22.8)	3.86 (0.45, 14.9)
ADA	4.9 (0.59, 18.7)		2.04 (0.09, 9.51)	2.99 (0.13, 13.9)	2.74 (0.05, 12.6)	1.14 (0.13, 4.37)
ETA	4.24 (0.72, 14.3)	2.06 (0.1, 9.48)		5.16 (0.09, 22.9)	2.53 (0.06, 11.5)	1.79 (0.08, 8.13)
IFX	10.6 (0.49, 48.9)	3.21 (0.15, 14.7)	5.19 (0.1, 23.7)		3.2 (0.02, 13.7)	0.76 (0.12, 2.6)
GOL	5.09 (0.31, 22.7)	2.64 (0.05, 12.6)	2.66 (0.06, 11.4)	3.27 (0.02, 12.6)		3.46 (0.06, 15.8)
RTX	3.85 (0.46, 14.8)	1.16 (0.14, 4.43)	1.78 (0.08, 8.04)	0.72 (0.12, 2.44)	3.35 (0.06, 15.4)	

This was due to the increase in the between-study heterogeneity, from between-study standard deviation of 0.7 (0.14, 1.81) from the NMA of registry data alone to 0.93 (0.37, 1.86) from NMA of the combined data for the adjusted results.

The results of a bivariate NMA combining data from RCTs and BSRBR-RA of biologic therapies in both lines of therapy are shown in Table 4, with the results using the “standard” bivariate NMA model in the upper triangle and the results from the analysis assuming exchangeability of the effects of the biologic therapies in the lower triangle. In both analyses, the adjusted data were used for the emulated target trials. Combining data from the first and second lines of therapy through the use of the bivariate NMA led to a decrease in uncertainty for many of the individual treatments when compared to the results of the univariate NMA of second-line therapy alone (combining RCT data with BSRBR-RA data). For example, the effect of etanercept vs. methotrexate on ACR20 response was OR = 3.8 (0.9, 11.1) from the bivariate NMA compared to OR = 4.3 (0.7, 14.4) from the univariate NMA. The between-studies correlation was weak, $\rho = -0.3$ (95% CrI: $-0.94, 0.66$), which limited the borrowing of information across the treatment lines. The heterogeneity parameter for the second-line treatments was 0.91 (0.36, 1.84), which was comparable with the results of the univariate NMA, and 0.93 (0.65, 1.35) for the first-line treatments.

The bivariate NMA approach assuming the additional exchangeability of the absolute effects of the biologic therapies allowed for predictions of treatment effects that had not been evaluated in trials in a second-line setting. In this case, it produced effectiveness estimates for abatacept in the second line of therapy against all other treatments in

the network. Moreover, this additional exchangeability led to a noticeable reduction in uncertainty around the remaining estimates of effect for other therapies, as can be seen in the lower triangle of Table 4. This was a result of additional borrowing of information across the biologic therapies. However, there may have been some degree of smoothing of the effects across the biologic therapies, which was difficult to judge due to the large uncertainty. A sensitivity analysis was carried out using a t -distribution in place of the normal distribution in model (3) in the Supplementary materials C, which largely produced very similar results but inflated the uncertainty around the effectiveness estimate for abatacept in the second line.

4. Discussion

There has been an increased interest in use of real-world evidence to inform clinical and policy decisions in healthcare. For example, Schünemann et al. discuss a range of scenarios where nonrandomized evidence can contribute as complementary, sequential, or replacement evidence for RCTs when evaluating the effectiveness of interventions in meta-analysis [19]. Our approach for combining evidence on early and late designs (first vs. second lines of therapy) provides a new highly informative way to use both sources of evidence in the networks. We provide a conceptual approach for using RWD, such as from registries or electronic health records, to generate estimates of effectiveness of treatments in first and second lines of therapy and combining them with RCT data to enhance the evidence base and provide effectiveness estimates of therapies in the second line, where data on effectiveness in the second line are not available from RCTs. In such

Table 4. Results of a bivariate NMA combining data from RCTs and BSRBR-RA of biologic in both lines of therapy using the “standard” bivariate NMA model (upper triangle) and assuming exchangeability of biologic therapies (lower triangle), represented as ORs with 95% CrIs

Treatment	MTX	ADA	ETA	IFX	GOL	ABT	RTX
MTX		4.1 (0.57, 15.3)	3.81 (0.87, 11.1)	7.64 (0.3, 36.2)	4.92 (0.32, 21.7)	–	4.4 (0.49, 17.3)
ADA	2.96 (1.51, 5.29)		2.01 (0.13, 8.64)	2.17 (0.15, 9.48)	2.94 (0.06, 13.4)	–	1.7 (0.13, 7.43)
ETA	3.07 (1.72, 5.18)	1.1 (0.58, 2.03)		3.92 (0.07, 16.7)	2.19 (0.07, 9.96)	–	1.91 (0.1, 8.69)
IFX	3.13 (1.38, 6.31)	1.09 (0.56, 2.03)	1.05 (0.47, 2.04)		4.79 (0.03, 20.0)	–	1.65 (0.09, 7.85)
GOL	3.03 (1.38, 5.78)	1.08 (0.49, 2.12)	1.02 (0.45, 1.91)	1.06 (0.42, 2.14)		–	4.03 (0.07, 18.3)
ABT	3.37 (1.08, 7.81)	1.19 (0.41, 2.76)	1.13 (0.36, 2.55)	1.16 (0.37, 2.64)	1.2 (0.38, 2.88)		–
RTX	3.09 (1.55, 5.83)	1.09 (0.57, 2.08)	1.04 (0.51, 1.97)	1.06 (0.52, 2.0)	1.11 (0.5, 2.31)	1.17 (0.39, 2.74)	

circumstances, producing these estimates would require conducting expensive and time-consuming additional clinical trials. The proposed approach can be used to carry out a feasibility analysis or provide inputs to the trial design or even be used for evidence-based decision-making where evidence is sufficiently robust.

When carrying out this research, we came across a number of limitations. Some of them were related to data. In particular, the RCT data were relatively sparse with a star-shaped network for the first-line treatments and only two trials reporting the effectiveness of biologic therapies in the second line. The dataset was simplified by combining the control arms (including methotrexate as either combination therapy or concomitant therapy with placebo) into the same control arm denoted as methotrexate. This was done to strengthen the network structure to better illustrate the methodological aspect of this work. Most of the biologic arms also included methotrexate. Considering that for a large proportion of visits in the BSRBR-RA data, methotrexate was recorded as a concomitant therapy to a biologic treatment; an assumption was made that a large proportion of patients receiving biologic therapy, across all studies, also received methotrexate. The registry dataset contained a substantial amount of missing data, in particular for some of the components of the ACR20 response criteria, which was not due to the issues of quality of the data but owing to the fact that some of the components are not routinely collected by the register. To estimate the response to the biologic therapies, we chose to relax the definition of the response. In addition, the register only captures 28 joint counts, which may be different from some of the trials. Considering these potentially strong assumptions around the data sources, the results of our analysis should not be used for clinical interpretation but only as an illustration of the proposed methodology.

There were only six target trials generated from the registry data, which resulted in substantial uncertainty around the between-studies correlation, as these were the only studies contributing data to estimating the correlation. The combined network was still limited with a lack of data on each contrast and line across study designs. Target trial data were incorporated in the NMAs at a face value, assuming they were equivalent with RCT data. Extensions of the analysis could include a power prior approach [20], allowing for down-weighting of RWD, or hierarchical modeling to differentiate between the two study designs [4]. Further investigation into data scenarios and model assumptions needs to be carried out to understand when this framework can be most efficient.

5. Conclusion

Registry data can be used to bridge networks of first and second lines of therapy which are disconnected when using RCT data alone. Bivariate NMA of combined data from RCTs and RWD can be used to predict effectiveness of a treatment in second-line use when the therapy is only

investigated in an RCT as first line (or vice versa). The approach can be applied to other settings where RCT data are available for disjoint subsets of population, such as, for example, children and adults and registries may provide data covering follow-up period from childhood to adulthood for each individual.

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Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.jclinepi.2022.06.011>.

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