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# Aggregation dynamics explain vegetation patch-size distributions

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## Abstract

Vegetation patch-size distributions have been an intense area of study for theoreticians and applied ecologists alike in recent years. Of particular interest is the seemingly ubiquitous nature of power-law patch-size distributions emerging in a number of diverse ecosystems. The leading explanation of the emergence of these power-laws is due to local facilitative mechanisms. There is also a common transition from power law to exponential distribution when a system is under global pressure, such as grazing or lack of rainfall. These phenomena require a simple mechanistic explanation. Here, we study vegetation patches from a spatially implicit, patch dynamic viewpoint. We show that under minimal assumptions a power-law patch-size distribution appears as a natural consequence of aggregation. A linear death term also leads to an exponential term in the distribution for any non-zero death rate. This work shows the origin of the breakdown of the power-law under increasing pressure and shows that in general, we expect to observe a power law with an exponential cutoff (rather than pure power laws). The estimated parameters of this distribution also provide insight into the underlying ecological mechanisms of aggregation and death.

*Keywords:* patch-size distribution, pattern formation, spatial ecology, aggregation

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## 1. Introduction

Vegetation patch-size distributions have been under intense study in recent years [1, 2, 3, 4, 5]. It has been shown that a power-law provides a good fit to the patch-size distribution under a robust range conditions, however there are marginal cases to this. Kéfi et al. [6] analysed patch-size distributions in semi-arid vegetation in the Mediterranean and found that there was not only a power-law distribution evident in the patch-size distribution, but also a truncated exponential term, when the system was under increased grazing pressure. Similar power-law distribution phenomena have also been detected in a number of other ecosystems including mussel beds [7] and marine benthic diatoms [8]. This phenomena of a power-law distribution transitioning to an exponential distribution under increasing stress has recently shown to be robust, where diverse ecological models are able to reproduce these results [2].

The leading explanation of this power-law pattern formation in ecology is due to local interactions driving the large-scale behaviour [9, 10]. Scanlon et al. [11] supported through the use of numerical simulation of spatially-explicit models of vegetation growth combined with a global effect on the population density interpreted as the amount of rainfall or other global processes. The local positive feedback process driving the patch formation is through facilitation of neighbourhood sites that increase the birth rate and decrease the death rate [5]. This explanation does not answer how a power-law forms at the patch level, whether it is due to a competition effect between larger clusters dominating the landscape or an aggregation of smaller clusters. There is also an open question of how patches aggregating together drives these observed patterns.

Models of aggregation and fragmentation have been considered in other areas in ecology such as the size of fish schools [12] and marine diatoms [13]. Aggregation phenomena has been more generally studied in the Physical sciences [14], including processes such as polymerisation [15], coagulation of aerosols [16] and flocculation [17]. Although these examples include clusters that may diffuse, aggregation phenomena may also be considered in the case where clusters are immobile [18]. Aggregation of vegetation clusters, however, has not been previously considered as an explicit driving force of the evolution of the patch-size distribution. Our novel contribution here is to apply established theory of aggregation dynamics to the system of vegetation clusters and derive a new model of aggregation with global death that

38 is applicable to vegetation dynamics.

39 In this article, spatially implicit models of vegetation clusters are investi-  
40 gated by considering how patches form and aggregate. The general conditions  
41 under which a power-law distribution is expected to emerge are explored as  
42 well as when there is a breakdown of the power law distribution due to an ex-  
43 ponential truncation. By adopting a patch-centric viewpoint, the impact of  
44 aggregation on the resulting distribution along with other processes may be  
45 studied directly. This represents a powerful new approach to understanding  
46 the origin of these distributions, by explicitly modelling the patch-size dy-  
47 namics without the need to infer the patch-size distribution from a spatially  
48 explicit model [5].

49 Further, the connection between the power-law exponent and the persis-  
50 tence of the distribution in this model is explored. We begin by defining a  
51 novel model of aggregation with linear death and then deriving an asymp-  
52 totic solution when the death rate is small. This analytic result is compared  
53 to a simulation study of vegetation with local and global growth properties  
54 subjected to a global disturbance. For small disturbance, the power law ex-  
55 ponent closely matches the exponent expected from the model. The conclu-  
56 sion is that the power-law clustering observed in many vegetation ecosystems  
57 may simply be an aggregation effect and the exponential truncation observed  
58 when there is increased stress is due to an increase in the linear death rate  
59 of clusters.

## 60 **2. Theory**

61 The idea developed here is to model the patches themselves as opposed  
62 to an individual spatial site as is done in probabilistic cellular automata  
63 [19, 20]. We denote  $c_k(t)$  as the density of patches of size  $k$  at time  $t$ , where  
64 time is taken to be continuous. A continuous model of patch-sizes can be  
65 studied, however for the present  $k$  shall take positive integer values only,  
66  $k \in \{1, 2, \dots\}$ . A kernel of aggregation gives the rate at which patches of  
67 size  $i$  and  $j$  aggregate together to form a patch of size  $i + j$ , this kernel  
68 is denoted  $K(i, j)$ . Finally it is assumed there is a constant rate at which  
69 patches of size 1 or monomers enter the system. These assumptions are  
70 general and can include many different phenomena, including static clusters  
71 and diffusing monomers [18]. The governing master equation, also known as

72 the Smoluchowski equation [21] is then

$$\frac{d}{dt}c_k = \frac{1}{2} \sum_{i+j=k} K(i, j)c_i c_j - \sum_{j \geq 1} K(j, k)c_j c_k + \delta_{k,1}, \quad (1)$$

73 where  $\delta_{k,1}$  is the Kronecker-delta function that is 1 when  $k = 1$  and 0 other-  
 74 wise. For convenience, time has been re-scaled such that the rate at which  
 75 aggregation occurs is 1. It is instructive to imagine a single unit or monomer  
 76 coming into contact with a cluster and calculating the rate at which this oc-  
 77 curs for larger as opposed to smaller clusters. If  $a > 0$  then, assuming the size  
 78 of the monomer is negligible, the monomer rate equation is  $K(i) = i^{-a}$ . This  
 79 means smaller clusters are favoured and the growth rate reduces as clusters  
 80 grow larger in size. An ecological explanation of this could be due to the  
 81 self-limitation through competition a larger cluster experiences with itself,  
 82 thus reducing its potential for growth. Smaller clusters have more space and  
 83 thus can grow at a quicker rate.

84 When  $a < 0$ , larger clusters are favoured for growth compared with  
 85 smaller clusters, this can be seen as a form of the Allee effect [22]. In the  
 86 regime when  $a < 0$ , small clusters are more susceptible to environmental  
 87 perturbation and as such, have a lower propensity for growth. At the other  
 88 length scale, larger clusters of vegetation are able to regulate their environ-  
 89 ment more and thus have greater resources for growth (An example species  
 90 where this holds is ribbed mussels [23], where larger clusters provide protec-  
 91 tion and shelter for new mussels). This example of an Allee effect can be  
 92 demonstrated by again considering the rate at which single units of vegetation  
 93 aggregate to a cluster. If a  $i > j$ , then  $K(1, i) = 1 + i^{-a} > 1 + j^{-a} = K(1, j)$   
 94 i.e. the rate at which a larger patch recruits new growth is greater than for a  
 95 smaller patch. A value for  $a$  then can give an indication of whether there is  
 96 strong small cluster growth at the expense of large clusters forming or if the  
 97 converse holds.

98 An alternative explanation of the aggregation exponent  $a$  is due to the  
 99 edge effects of a cluster. A single individual vegetation unit aggregates to a  
 100 cluster proportional to the edge of that cluster. If all clusters are non-fractal  
 101 then it would be expected that a vegetation unit aggregates at rate  $i^{1/2}$ , since  
 102 the length of a non-fractal object scales as a square root with its area. For  
 103 a general fractal cluster with boundary dimension  $d$ , it would be expected  
 104 that an individual unit scales as  $i^{1/d}$ .

105 Various properties are desirable for the kernel. Firstly symmetry, where  
 106 the rate at which patches of size  $i$  and  $j$  aggregate does not depend on the

107 ordering of the patches i.e.  $K(i, j) = K(j, i)$ . Secondly, scaling homogeneity,  
 108 where the rate at which patches of a certain size aggregate scales by some  
 109 factor  $K(mi, mj) = m^\lambda K(i, j)$ . The simplest kernel that satisfies these con-  
 110 ditions is the constant kernel  $K(i, j) = 1$ , corresponding to the case where  
 111  $\lambda = 0$ . When this form of kernel is assumed, the tail-solution (for large  $k$ )  
 112 has the simple form [24]

$$c_k \sim \frac{1}{\sqrt{4\pi}} \frac{1}{k^{3/2}}. \quad (2)$$

113 The tail of the patch-size distribution is a power law with exponent  $3/2$ ,  
 114 where the power law nature of the solution is a consequence of the injection  
 115 term ( where births of patch size one enter the system ) and the non-linear  
 116 aggregation term in the equation. The equation can be solved analytically  
 117 for more general kernels of the type

$$K(i, j) = i^{-a} + j^{-a}. \quad (3)$$

119 This type of kernel also admits an analytic solution in the large patch-size  
 120 limit [25, 26] with a steady state distribution of the form where

$$c_k \sim Ck^{-\tau} \quad (4a)$$

$$\tau = \frac{3-a}{2}, \quad C = \sqrt{\frac{1-a^2}{4\pi} \cos\left(\frac{\pi a}{2}\right)}. \quad (4b)$$

121 For a steady state to exist we require  $-1 < a < 1$  and hence the scaling  
 122 exponent can be found on the interval  $\tau \in (1, 2)$ . The dynamics of the  
 123 equation can be assessed by defining the cross-over time, which is the time  
 124 taken for a density of patches of a certain size to reach its asymptotic value.  
 125 The cross-over time for a patch of size  $k_*$  to the steady state solution  $c_{k_*}$   
 126 is found to take the form  $t = (k_*)^z$  where  $z = (1+a)/2$ . The scaling of  
 127 the cross-over time and the patch-size exponent can be related by the simple  
 128 linear equation  $\tau = 2 - z$ . This gives a linear relationship between the static  
 129 exponent at stationarity and its dynamic exponent.

130 A real vegetation system is not purely defined by an aggregation process  
 131 however. In particular in the previous example there is no death either of  
 132 single vegetation units or patch clusters. Death may lead to changes in the  
 133 exponent of the stationary distribution and so it is important to include in  
 134 any model of vegetation clustering. It is also assumed that a death event does  
 135 not lead to fragmentation of the cluster. A modified Smoluchowski equation

136 with a linear death term can then be produced as

$$\frac{d}{dt}c_k = \frac{1}{2} \sum_{i+j=k} K(i, j)c_i c_j - \sum_{j \geq 1} K(j, k)c_j c_k + \mu(k+1)c_{k+1} - \mu(k)c_k, \quad (5a)$$

$$\frac{d}{dt}c_1 = - \sum_{j \geq 1} K(j, 1)c_j c_1 + 1 + \mu(2)c_2 - \mu(1)c_1. \quad (5b)$$

137 The general additive aggregation kernel is again taken to be of the form  
 138  $K(i, j) = i^{-a} + j^{-a}$ , where  $a$  represents the scaling parameter of the rate at  
 139 which aggregates of a certain size join. If it is equally likely for a cluster  
 140 of a certain size to aggregate with a cluster of any other size then the scale  
 141 parameter  $a = 0$ . For a pure aggregation system with no fragmentation, this  
 142 leads to a cluster scaling of  $3/2$ .  $\mu(k)$  defines the death rate, which is the  
 143 rate at which individual units are lost from a patch, where a patch of size  $k$   
 144 transitions to a patch of size  $k - 1$  due to exogenous or endogenous factors.  
 145 A number of different forms of this death rate may be considered dependent  
 146 on the biological details of the system. For example if each individual has a  
 147 constant rate of death regardless of the size of patch its contained, such as due  
 148 to lack of rainfall or grazing, is then  $\mu(k) = \mu k$ . If death occurs at the edge  
 149 of a patch then the death rate is  $\mu k^{1/d}$ , where  $d$  is the boundary dimension  
 150 of the patches. The simplest form of the death rate is where  $\mu(k) = \mu$  for  
 151 all  $k$ . In order to gain insight into the effect of a death rate on the resulting  
 152 patch-size distribution, we assume the final form of the death rate.

153 In order to gain analytic tractability on the model a constant aggregation  
 154 kernel is assumed ( $a = 0, K = 2$ ) together with a constant death rate for each  
 155 individual within a patch. The strategy for deriving a solution is similar to  
 156 the strategy in Krapivsky et al. [26]. A constant kernel  $K(i, j) = 2$  is used.  
 157 Eq. (5) is rewritten as

$$\frac{d}{dt}c_k = \sum_{i+j=k} c_i c_j - 2c_k \sum_{j \geq 1} c_j + \mu c_{k+1} - \mu c_k, \quad (6a)$$

$$\frac{d}{dt}c_1 = -2c_1 \sum_{j \geq 1} c_j + 1 + \mu c_2 - \mu c_1. \quad (6b)$$

158 The asymptotic tail of the resulting patch-size distribution is then sought  
 159 in order to gain an understanding of how the linear death rate affects the  
 160 stationary distribution. By using the asymptotic approximation and assum-  
 161 ing  $k$  is large, the  $k$ -th coefficient in this expansion and hence the density of

162 patches of size  $k$  is

$$c_k = k^{-3/2} \exp(-\Lambda k), \quad (7)$$

163 where  $\Lambda = \log(1 + \mu N)$  and  $N$  is the total population size (See Appendix A  
164 for a derivation). The solution is therefore a power law with an exponential  
165 truncation of factor  $\Lambda$ . When the death rate is 0,  $\Lambda = 0$  and hence the patch-  
166 size distribution is a pure power law as is expected. A large death rate will  
167 lead to a solution that is dominated by an exponential decay term, hence the  
168 patch-size distribution is expected to have a smooth transition from a pure  
169 power law to an exponential distribution. A dimensionality argument of Eq. 6  
170 [27] also leads to a power law exponent of the form  $3/2$ . This solution can  
171 be compared to the general power-size distribution with exponential cut-off  
172  $N(k)$  given by

$$N(k) = Ck^{-\alpha} \exp(-k/k_x), \quad (8)$$

174 where  $k_x$  is the patch-size above which  $N(s)$  decreases faster than power-law  
175 [2]. Matching terms and assuming  $\mu N$  is small gives the following simple  
176 relationship between the cross-over patch size  $k_x$  and death rate  $\mu$  as

$$k_x = \frac{1}{N\mu}. \quad (9)$$

178 The model therefore predicts an inverse relationship between patch-size cross-  
179 over and death rate. This also predicts that when the death rate is small  
180 enough, the cross-over patch-size  $k_x$  may be larger than the system size and  
181 as such the exponential tail may not be observed in empirical distributions.

### 182 3. Results

183 In order to compare the model predictions of patch formation in an aggre-  
184 gation system with a constant death rate to the prediction of the patch-size  
185 distribution obtained in Eq. 7 is compared to a simple probabilistic cellu-  
186 lar automata model of vegetation growth. The cellular model is similar to  
187 the one discussed in [11], the model is defined on a toroidal lattice where  
188 each site can exist in one of two states: occupied (1) and empty (0). The  
189 occupied state propagates through nearest neighbour growth at rate  $\beta$ , as  
190 well as through a background constant birth probability  $\gamma$ . The alive sites  
191 transition to a dead site with a constant death probability  $\mu$ . Hence if  $n_x$   
192 is the number of alive sites in the neighbourhood of site  $x$ , the transitions can

193 be summarised as

$$P_x(0 \rightarrow 1) = \epsilon \min\{1, \gamma + \beta n_x/4\}, \quad (10a)$$

$$P_x(1 \rightarrow 0) = \epsilon\mu, \quad (10b)$$

194 where  $\epsilon$  sets the total reaction rate of the system and was implemented to  
195 reduce the probability of multiple events occurring within the same neigh-  
196 bourhood. The minimum function is used here to guarantee the probability  
197 of transitioning to an alive state is one in the rare case when the sum of the  
198 two probabilities increases above one.

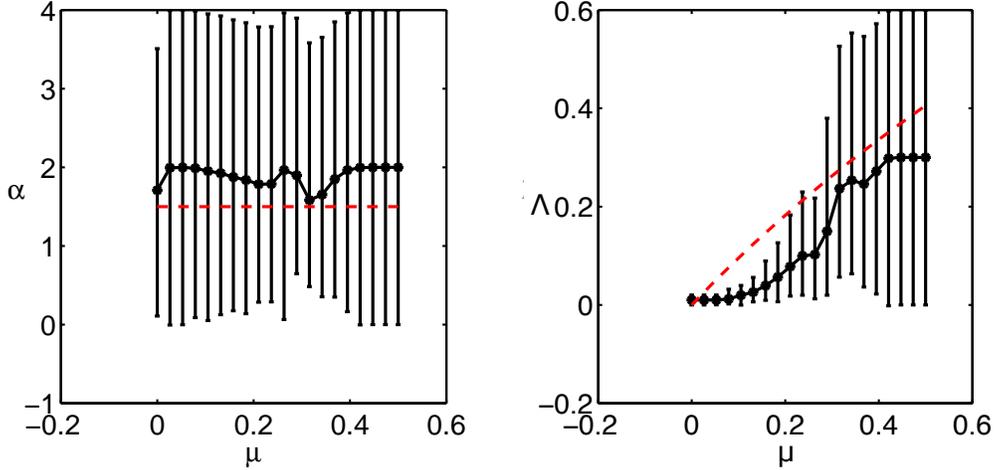
199 Simulations were conducted for constant local growth, birth rate and  
200 reaction rate  $\beta = 0.2, \gamma = 0.01, \epsilon = 0.1$  and over a range of death rates.  
201 Simulations were ran for lattice length  $L = 500$  and for 1000 replicates of each  
202 parameter set. The patch-size distribution was recorded for each simulation  
203 run after 600 time-steps. This was chosen so that when  $\mu = 0$ , the lattice  
204 is approximately 50% occupied. The following power-law with exponential  
205 truncation was fitted to the distribution using a maximum likelihood method

$$f(K = k) = Ck^{-\alpha} \exp(-\Lambda k), \quad (11)$$

206 for some normalising factor  $C$ . The resulting maximum likelihood estimators  
207 were found using a downhill simplex method implemented in Matlab  
208 R2014a [28]. The approximate solution to the aggregation equation predicts  
209 a constant power-law exponent  $\alpha$  of  $3/2$ . This is close to the inferred value  
210 from simulation for the range of  $\mu$  values studied (Fig. 1a). The exponential  
211 factor  $\Lambda$  is zero when the death rate is zero (Fig. 1b), as predicted. For  
212 increasing death rate,  $\Lambda$  increases again as predicted. Overall there is an  
213 increase in the exponential factor for increasing death rate as is predicted,  
214 however the functional form of the increase is not fully captured by the mean  
215 field approximation.

#### 216 4. Discussion

217 Changing the focus away from explicit spatial modelling of vegetation  
218 patch formation and instead focusing on the dynamics of patch-sizes gives a  
219 unique insight into the underlying aggregation-fragmentation processes. Here  
220 we have primarily focused on solutions to equations where the aggregation  
221 kernel that governs the rate at which patches of two sizes will aggregate by



(a) comparison of power-law exponent from simulation to theoretical prediction (b) comparison of exponential factor from simulation to theoretical prediction

Figure 1: Exponents of patch-size distribution compared to simulations. Theoretical values shown as red dashed lines, while simulation calculated values are given as black dots with 95% confidence intervals. The theoretical values for the power-law exponent  $\alpha$  and the exponential factor  $\Lambda$  are derived in Eq. 7. As predicted for small values of the death rate the power-law component of the patch-size distribution is constant whilst there is an increase in the exponential component for increasing death rate.

222 either a constant or power law kernel. For a system where there is aggrega-  
 223 tion only the resulting patch-size distribution is that of a pure power law,  
 224 with exponent that is dependent on the exponent of the power law aggrega-  
 225 tion kernel. The introduction of a linear death term, where an individual is  
 226 lost from a patch at rate  $\mu/k$  gives rise to a power law with exponential tail  
 227 distribution of the form  $c_k \sim k^{-\alpha} \exp(-\Lambda k)$ . This solution holds generally  
 228 when there is a linear death term and power-law aggregation kernel, even  
 229 when the kernel is composed of a sum of two power-laws. Further,  $\alpha$  is de-  
 230 pendent on the specifics of the aggregation term alone and  $\Lambda$  is dependent on  
 231 the death rate alone. This separation of the aggregation and fragmentation  
 232 term implies, in principle, the ability to infer aggregation and death processes  
 233 through observing the converged patch-size distribution alone, hence this is  
 234 applicable to inferring process from a single spatial snapshot.

235 Kéfi et al. [6, 2] predicts that a power-law distribution in the patch-size  
236 distribution occurs when a global environmental death rate is small. This  
237 transitions to an exponential distribution when there is greater stress on the  
238 system through this global death rate term. The model used is a spatially-  
239 explicit one with a local growth term and a background death rate. The  
240 model proposed in this article can be seen as a deterministic equivalent of  
241 this spatially-explicit model. Through the derived solution in this article it  
242 is observed that there should always be an exponential tail to the distribution  
243 if the death rate is non-zero. Similar arguments have also been made recently  
244 [29], but notably none have explained the origin of a power law with exponen-  
245 tial tail observed in vegetation systems. The derived model then, provides a  
246 theoretical origin to the observed spatial patterns in vegetation ecosystems  
247 that are under a pressure that can be considered constant throughout space  
248 (rainfall, grazing etc.). As an example, if all sites had the same death rate  
249 regardless of its neighbourhood, such as for a grazer, the death rate would  
250 be  $\mu k$  for a patch of size  $k$ . The model also suggests that a power-law with  
251 exponential tail is a more accurate description of the patch-size distribution,  
252 although when the death rate is small, the exponential tail may not be ob-  
253 served directly. This approach would be able to provide further insights into  
254 the nature of the patch-size distribution for other systems where disturbance  
255 may be spatially distributed.

256 The model also gives insight into how there can be a continuous array  
257 of power-law exponent observed in nature. The aggregation with no death  
258 model predicts that power-laws exponents in the range  $(1, 2)$  are physically  
259 possible, which is what has been observed in a number of ecosystems [7, 6, 8].  
260 The model predicts that a change in the exponent of a patch-size distribu-  
261 tion is related to a change in the power-law exponent of the aggregation  
262 kernel. A simple dimensionality argument can be used to show that in the  
263 aggregation and death model with a kernel that has a general power law  
264 scaling as described in Eq. 3, the resulting stationary distribution will have  
265 the same exponent as that in the model with no death [27]. The conclusion  
266 of how to relate the patch-size distribution to the system dynamics is that  
267 both the power-law exponent and the presence of an exponential cut-off does  
268 give an indication of the underlying dynamics. More complex fragmentation  
269 processes than the one discussed would alter these conclusions however, as a  
270 non-linear fragmentation process will also lead to self-similar solutions and  
271 thus the two processes are confounded when only the stationary state is ob-  
272 served [30], such processes include storms and other strong weather events

273 that could split a single cluster of vegetation into multiple clusters. The size  
 274 of the system where the dynamics occur, such as in the lattice model, may  
 275 also have an impact on the exponents of the patch-size distribution due to  
 276 finite-size effects [31].

277 Other possible extensions to the model could include multi-species sys-  
 278 tems, where patches are formed of multiple species each with their own in-  
 279 trinsic death rates. Multi-species systems have already been considered in  
 280 the physical sciences and as such this would make for an interesting avenue of  
 281 future research [32]. Where the aggregation process is indistinguishable be-  
 282 tween two different species, this leads to similar results laid out in this article  
 283 [33]. However, more complex interactions such as inter-specific competition  
 284 would inevitably lead to a more complex relationship between the exponent  
 285 death term and the underlying death rates. The model equations were scaled  
 286 such that the rate of aggregation and rate at which single vegetation units  
 287 are created is one. This was done for convenience since we were interested in  
 288 studying the scaling alone, whereas these parameters change the constant of  
 289 the patch-size distribution only. Another extension then would be to explic-  
 290 itly calculate the constant for the patch-size distribution and study how this  
 291 changes as the other system rates change.

## 292 Acknowledgement

293 The work conducted in this article was funded as part of an EPSRC  
 294 studentship.

## 295 Appendix A. Derivation of asymptotic solution

296 A moment-generating function is used to find the steady state solution  
 297 to this equation in a similar fashion to the one described in Krapivsky et al.  
 298 [26]. Firstly define the total number of all patches as  $N = \sum_{k \geq 1} c_k$  and then  
 299 sum Eq. (6a-b) in order to obtain

$$\frac{dN}{dt} = \sum_{k \geq 1} \sum_{i+j=k} c_i c_j - 2 \sum_{k \geq 1} c_k \sum_{j \geq 1} c_j + 1 + \sum_{k \geq 1} \mu c_{k+1} - \sum_{k \geq 1} \mu c_k, \quad (\text{A.1})$$

$$\frac{dN}{dt} = N^2 - 2N^2 + 1 - \mu c_1, \quad (\text{A.2})$$

$$\frac{dN}{dt} = -N^2 + 1 - \mu c_1. \quad (\text{A.3})$$

300 Dynamically, consider when  $N$  is at equilibrium. If  $\mu = 0$  then the stationary  
 301 solution is  $N = 1$ . If  $\mu > 0$  then the equilibrium solution is necessarily  
 302 bounded between one and zero as  $N$  and  $c_1$  are always positive.

303 The moment-generating function  $C(z, t) = \sum_{k=1}^{\infty} c_k z^k$  is now considered.  
 304 Multiplying Eq. (6) by  $z^k$  and summing over all  $k$  gives the following

$$\begin{aligned} \frac{d}{dt}C &= C^2 - 2NC + z + \mu \sum_{k \geq 1} z^k c_{k+1} - \mu \sum_{k \geq 1} z^k c_k \\ &= C^2 - 2NC + z + \frac{\mu}{z}C - \mu C - \mu c_1. \end{aligned} \quad (\text{A.4})$$

305 The  $C^2$  term is derived using the relationship

$$\left( \sum_{k \geq 1} a_k \right)^2 = \left( \sum_{i \geq 1} a_i \right) \left( \sum_{j \geq 1} a_j \right) = \sum_{k \geq 1} \sum_{i+j=k} a_i a_j. \quad (\text{A.5})$$

307 The new moment generating function defined as  $A(z, t) = C(z, t) - N(z, t)$  is  
 308 considered in order to derive the final stationary solution. The time derivative  
 309 is calculated by combining Eq. A.4 with Eq. A.3

$$\begin{aligned} \frac{d}{dt}A(z, t) &= \frac{d}{dt}C(z, t) - \frac{d}{dt}N(t) \\ &= C^2 - 2NC + \frac{\mu}{z}C + z - \mu C - \mu c_1 - 1 + N^2 + \mu c_1 \\ &= A^2 + \frac{\mu}{z}C - \mu C + z - 1 \\ &= A^2 + \mu \frac{1-z}{z}A + \mu \frac{1-z}{z}N + z - 1. \end{aligned} \quad (\text{A.6})$$

310 Note that the right-hand side is quadratic in terms of  $A$ . Setting the time-  
 311 derivative to zero gives the steady-state solution of the moment-generating  
 312 function as

$$A = \mu \frac{z-1}{z} + \sqrt{\mu^2 \frac{(1-z)^2}{z^2} - 4 \left( \mu \frac{1-z}{z}N + z - 1 \right)}. \quad (\text{A.7})$$

313 In order to proceed it is assumed that the death rate  $\mu$  is small and only the  
 314 leading order term is kept. Hence

$$A \approx 2\sqrt{1-z - \mu \frac{1-z}{z}N}. \quad (\text{A.8})$$

315 The strategy is to find  $A$  in terms of the power series  $\sum_{k=1}^{\infty} c_k z^k$ , where  $c_k$  is  
 316 a function of  $\mu$ . Assuming  $z$  is sufficiently close to one such that  $z + \mu(1 -$   
 317  $z)N/z < 1$ , the expansion of  $\sqrt{1-x}$  is used to obtain

$$A_{\text{approx}} = 2 \sum_{k=0}^{\infty} \frac{\Gamma(3/2)}{\Gamma(3/2 - k)\Gamma(k+1)} (1 + \mu N)^{1/2-k} (-z - \mu N/z)^k. \quad (\text{A.9})$$

318 Using the relationship  $\Gamma(z)\Gamma(1-z) = \frac{\pi}{\sin(\pi z)}$ , cancelling the  $(-1)^k$  terms and  
 319 absorbing all constants into a constant  $c$  term

$$A_{\text{approx}} = c \sum_{k=0}^{\infty} \frac{\Gamma(k-1/2)}{\Gamma(k+1)} (1 + \mu N)^{1/2-k} (z + \mu N/z)^k. \quad (\text{A.10})$$

320 Using the binomial expansion, this becomes

$$A_{\text{approx}} = c \sum_{k=0}^{\infty} \sum_{i=0}^k \frac{\Gamma(k-1/2)}{\Gamma(k+1)} \frac{\Gamma(k+1)}{\Gamma(i+1)\Gamma(k-i+1)} (1 + \mu N)^{1/2-k} (\mu N)^{k-i} z^{2i-k}. \quad (\text{A.11})$$

321  
 322 In order to find the  $k$ -th coefficient as  $k \gg 1$  the leading order of the  $i$  term  
 323 in the binomial is considered. Given that  $\mu N \ll 1$ , the  $i$  dependent terms  
 324 are dominated by  $i = k$ . Hence the  $k$ th term of the expansion where  $k$  is  
 325 large is

$$326 \quad c \frac{\Gamma(k-1/2)}{\Gamma(k+1)} (1 + \mu N)^{1/2-k} z^k. \quad (\text{A.12})$$

327 By using the asymptotic approximation  $\Gamma(n+a)/\Gamma(n) \sim n^a$  and assuming  
 328  $k$  is large, the  $k$ -th coefficient in this expansion and hence the density of  
 329 patches of size  $k$  is

$$330 \quad c_k = ck^{-3/2} \exp(-\Lambda k), \quad (\text{A.13})$$

331 where  $\Lambda = \log(1 + \mu N)$ .

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