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1	Design resistance of helical seam pipe columns with limited
2	tensile test data
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# 11 ABSTRACT

This study conducts a capacity factor calibration for steel and steel-concrete composite columns with 12 helical seam pipe (also known as spiral welded tube) sections considering the following three member 13 types: (i) steel columns under axial compression, (ii) concrete-filled steel tubes (CFSTs) under axial 14 compression, and (iii) CFSTs under combined axial compression and uniaxial bending due to load 15 eccentricity. The calibration has been conducted for both forward and inverse reliability analyses. 16 The forward analysis calibrates the capacity factor of steel contribution in the design models given in 17 NZS 3404.1, AS 4100, AS/NZS 5100.6 and AS/NZS 2327 to meet the target reliability level provided 18 19 in both ISO 2394 and AS 5104 when using API 5L products in non-composite and composite column applications. Whilst the inverse analysis estimates the required minimum number of steel tensile tests 20 when the target reliability level and the capacity factors are all provided. In these analyses, a total of 21 22 68 experimental data collected from the literature are utilised to estimate the modelling uncertainty in terms of bias and scatter. For all member types, the design models achieve similar or higher 23 reliability than the target reliability, and the corresponding required minimum amount of steel 24 25 material tests are calculated and provided.

- 26
- *Keywords:* API 5L; steel-concrete composite columns; steel columns; design resistance; capacity
  factors; steel reuse.
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# 32 **1. Introduction**

## 33 1.1 Background

Steel and concrete composite construction has achieved a high market share in multi-storey buildings 34 35 and bridges. Composite columns, or concrete-filled steel tubes (CFSTs), using circular hollow sections have become popular because: greater resistance than rectangular or square hollow sections 36 is achieved through confinement of the concrete core [1,2,3]; the rules in current design standards 37 38 such as AS/NZS 2327 [4], EN 1994-1-1 [5], AISC 360-16 [6] imply that more structurally efficient cross-sections can be achieved, due to reduced local bucking from the presence of the concrete infill 39 (however, it has recently been shown that this enhancement may be limited [7]); and improved 40 ductility and damping characteristics [8]. The presence of the steel section also eliminates the need 41 for formwork and, as the erection schedule is not dependent on concrete curing time in these situations, 42 CFSTs improve the speed of construction [9]. 43

Whilst there has been a focus on the performance of CFSTs using standard section sizes, tailor-made circular sections are becoming popular with designers as larger diameter (or non-standard diameter) tubes can be specified in order that more structurally efficient solutions can be achieved. One of the most popular types of tailor-made circular section are helical seam pipe, which are also known as spiral welded tubes (SWTs). SWT's have historically been used in the conveyance of liquids or gases,

and are fabricated by helically bending a continuous length of steel plate and welding the abutting 49 edges [10]. The benefit of manufacturing SWTs compared to longitudinally welded tubes (LWTs) 50 51 arise from the efficiencies generated from the continuous spiral welding process, the lower cost of SWT manufacturers, and the fact that SWTs of different diameters can be readily manufactured using 52 the same forming tools. One of the most widely used product standards internationally for SWTs is 53 API 5L [11]. This is recognized by the International Organization for Standardization (ISO) where 54 continued harmonisation of ISO 3183 [12] with API 5L has led to the 2019 version of this standard 55 only presenting supplementary rules to complement API 5L, thereby resulting in a document that is 56 57 around six-times shorter than previous versions of ISO 3183.

In the construction sector, API 5L products have been successfully used for columns in multi-storey 58 buildings in the Asia-Pacific region for several years. Based on this long experience, the Hong Kong 59 Buildings Department [13] and the Building Construction Authority in Singapore [14] permit the use 60 of API 5L products in the design of steel structures. Whilst full-scale tests on both non-composite and 61 composite columns using SWTs have been conducted, and the results compared with the predictions 62 given by different design standards, the present authors are unaware of any structural reliability 63 analyses that have been undertaken to evaluate the capacity reduction factors required for design. 64 Moreover, given that API 5L SWTs are normally used for conveyance of liquids or gasses, the yield 65 strength of the material is measured in the circumferential direction. As a consequence of this, there 66 is a need to define the required minimum yield strength of the steel in the longitudinal direction when 67 SWTs are designed as columns. 68

69 The magnitude of the capacity factors given in steel and composite design standards are often based 70 on production control measurements of geometry and yield strength from steel producers that are

servicing a particular market at that time [15]. However, if different steel producers subsequently 71 enter the market, this could potentially invalidate the magnitude of the capacity factors previously 72 73 calculated because the required statistical input values are not normally given in product standards [16]. This issue was recently identified in the RFCS SAFEBRICTILE project, where it was intended 74 that the statistical properties evaluated for design could be included within future versions of the 75 European product standards [17]. Moreover, there is also a growing need to reuse reclaimed steel in 76 structural applications due to its environmental advantages. In these circumstances, not all of the 77 material and geometrical characteristics are known. Recent design guidance has recommended that 78 more onerous capacity factors than currently given in design standards should be used [18], but these 79 are not based on structural reliability analyses. Conversely, when tensile coupons are taken from the 80 reclaimed steel products and tests are undertaken, a calculation procedure is given to evaluate the 81 characteristic mechanical properties from the results [18], but what impact this has on design 82 resistance of the composite member remains unknown. 83

In the present paper, the sensitivity of the capacity reduction factors for the design models given in 84 NZS 3404.1 [19], AS 4100 [20], AS/NZS 5100.6 [21] and AS/NZS 2327 [4] are investigated when 85 using API 5L [11] products in non-composite and composite column applications. Whilst the focus 86 of this paper is to support SWTs in Australian/New Zealand standards, given that the design models 87 are similar to those used in other international standards, coupled with the fact that the reliability 88 analyses are conducted according to ISO 2394 [22] (in Australia and New Zealand AS 5104 [23] is 89 an identical adoption of ISO 2394), the results from this paper have a wider international relevance. 90 In a similar way as an earlier study that considered the reliability of beams and columns using standard 91 section sizes [24], it is proposed to investigate the sensitivity of using API 5L products by undertaking 92 inverse reliability analyses for each of the design models to identify the minimum number of 93

94 mechanical tests that would be required to deliver a particular capacity reduction factor. This latter
95 information will be of particular importance to designers when the nominal yield strength of a product
96 is unknown.

## 97 1.2 Research objectives

In this study, the following research objectives are identified: (i) experimental databases for steel and steel-concrete composite columns with SWTs are established by collecting data from the literature. Then, reliability analyses are conducted to evaluate the capacity reduction factors for the API 5L [11] steel grades identified below using the design models given in NZS 3404.1 [19], AS 4100 [20], AS/NZS 5100.6 [21] and AS/NZS 2327 [4]. (ii) Inverse reliability analyses are undertaken for each of the design models to identify the minimum number of mechanical tests that will be required to deliver the target capacity reduction factors.

Whilst API 5L [11] provides a wide range of sections, from consultation with the industry it was proposed to narrow the scope to include only welded steel (seamed) pipe, considering the cost premium for seamless pipes. On this basis, the study was confined to product specification level 2 (PSL 2) pipes which have a maximum  $f_y/f_u$  ratio of 0.93, using the following steel grades: X42 (290), X46 (320), X52 (360), X56 (390), X60 (415), X65 (450), X70 (485), and X80 (555) [N.B. The values in the parenthesis indicate the nominal yield strengths in MPa].

# 111 **2. Reliability based capacity factor calibration**

## 112 **2.1 Target reliability level**

113 Capacity factors are required to be calibrated to meet the acceptable level of consequences of 114 structural design failure. This acceptable level is described by a target failure probability or a 115 reliability index. The failure probability  $P_{\rm f}$  and the reliability index  $\beta$  has the following relationship:

$$P_f = \Phi(-\beta)$$

116 where  $\Phi$  is the cumulative distribution function (CDF) of the standardised normal distribution.

Both ISO 2394 [22] and AS 5104 [23] suggest that, at the ULS when the failure costs are large and the relative costs of safety measures are normal, the target reliability index related to a one-year reference period is  $\beta_1 = 4.4$ . However, in most design standards around the world the evaluation of the capacity reduction factors have been based on a 50-year reference period where the previous version of ISO 2394 [25] and EN 1990 [26] give  $\beta_{50} = 3.8$ . According to EN 1990, the values of  $\beta$ for a different reference period can be calculated using the following equation:

$$\Phi(\beta_n) = \left[\Phi(\beta_1)\right]^n \tag{2}$$

where  $\beta_n$  is the reliability index for a reference period of *n* years and  $\beta_1$  is the reliability index for one-year.

From using Eq. (2) it can be seen that for  $\beta_{50} = 3.8$  this is equivalent to  $\beta_1 = 4.7$ . However, considering that full mutual independence of failure events in subsequent years and expressed as the multiplications of the terms. This assumption is unrealistic, and it is suggested that  $\beta_{50} = 3.8$  can be more realistically interpreted to be  $\beta_1=4.4$  [27]. From this finding, and for consistency with the basis of international design standards for steel and composite construction, the target reliability index used in the present paper is based on a 50-year reference period with  $\beta_t = \beta_{50} = 3.8$ .

When we consider a case that the resistance has an unfavourable value without considering the loadeffect, this probability has the following relationship with the corresponding reliability index:

$$P(R \le R_d) = \Phi(-\alpha_d \beta_t) = \Phi(-\beta_R) \tag{3}$$

where  $\alpha_d$  is the design value for the first-order reliability method (FORM) [28] influence coefficient which has the value of 0.8 according to ISO 2394 [22]/AS 5104 [23], *R* is the resistance, and *R<sub>d</sub>* is the design resistance.

# 136 2.2 Capacity factor calibration procedure considering the effect of statistical uncertainty in 137 material tests

The capacity factor calibration procedure proposed in Kang et al. [24] based on EN 1990 Annex D.8 138 [26] is used to carry out the following two types of analyses by considering the effect of statistical 139 uncertainty in material tests: (i) forward analysis for calibrating a capacity factor for steel contribution 140 in a structural member; and (ii) inverse analysis to estimate the minimum number of tensile tests for 141 the steel yield strength. This method is selected because it considers that the resistance value always 142 has a non-negative value, and thus it follows a lognormal distribution. This method calibrates capacity 143 factors separately from load factors using the concept of the first-order reliability method (FORM) 144 [27], where sensitivity factors separate the calibration procedures for resistance and load. In this 145 method, the modelling error of the prediction models is statistically estimated from the comparison 146 between the model predictions and the experimental observations for those predictions. 147

148 In the forward analysis, the capacity factor ( $\phi$ ) for a resistance prediction model is defined as follows:

$$\phi = \frac{R_d}{R_n} \tag{3}$$

149 where  $R_d$  = the design resistance to meet the target reliability for resistance, and  $R_n$  = the nominal 150 resistance.

The calculation of these resistance parameters are carried out as follows: let us assume that  $g_R(\mathbf{x})$  is a resistance prediction model and  $\mathbf{x}$  is a vector of mean-measured values for design parameters. The 153 constant bias of this model can be statistically represented as follows:

$$\overline{b} = \frac{1}{N} \sum_{i=1}^{N} \left( \frac{R_{ei}}{R_{ii}} \right)$$
(4)

where N = the number of experimental data,  $R_{ei}$  = the resistance observed from the *i*<sup>th</sup> specimen in the experiment, and  $R_{ti}$  = the resistance predicted by the resistance prediction model  $g_R(\mathbf{x}_i)$  for the *i*<sup>th</sup> specimen. The unbiased resistance prediction R using this bias correction term is calculated as follows:

$$R = bg_R(\mathbf{x})\delta \tag{5}$$

157 where  $\delta$  = the modelling error of the unbiased resistance prediction. The modelling error for the *i*<sup>th</sup> 158 experiment,  $\delta_i$ , can be estimated as follows:

$$\delta_i = \frac{R_{ei}}{\overline{b}R_{ii}} \tag{6}$$

As it is assumed that R in Eq. (5) follows a lognormal distribution with non-negative values, the coefficient of variation (COV) of R is estimated as follows:

$$V_R \cong \sqrt{\left(V_\delta^2 + V_{Rt,inf}^2 + V_{Rt,finite}^2\right)} \tag{7}$$

where  $V_{\delta}$  = the COV of the modelling error estimated using Eq. (6), and  $V_{Rt,inf}$  = the COV of the parametric uncertainty for the parameters with an infinite number of experiments, and  $V_{Rt,finite}$  = the COV of the parametric uncertainty for the parameters with a finite number of experiments. This equation assumes that the modelling error and the parametric uncertainty are statistically independent.  $V_{Rt}$  can be calculated by using the Monte Carlo simulations or the first-order approximations. The standard deviation of  $\ln R$  ( $\sigma_{\ln R}$ ) is estimated as follows:

$$\sigma_{\ln R} = \sqrt{\ln\left(1 + V_R^2\right)} \tag{8}$$

167 which is used to calculate the design resistance  $(R_d)$  in Eq. (3) as follows:

$$R_{d} = \overline{b}g_{R}(\mathbf{x})\exp\left(-k\sigma_{\ln R} - 0.5\sigma_{\ln R}^{2}\right)$$
(9)

168 where

$$k = \frac{\left(k_{d,m}V_{\delta}^{2} + \beta_{R}V_{Rt,inf}^{2} + k_{d,p}V_{Rt,finite}^{2}\right)}{V_{R}^{2}}$$
(10)

169 where  $k_{d,m}$  = the fractile factor determined for a finite number of structural member tests, and  $k_{d,p}$ 170 = the fractile factor determined for a finite number of structural material tests, to represent the 171 statistical uncertainty due to the finite number of test data. These factors are for the target reliability 172 index for resistance ( $\beta_R$ ) at the 75% confidence level and can be calculated for unknown  $\sigma_{\ln R}$  as 173 follows:

$$k_d = t_\beta (n-1) \times (1+1/n)^{0.5}$$
(11)

where  $t_{\beta}(n-1)$  is the fractile of the *t*-distribution for the probability corresponding to  $\beta_R$ . This fractile factor is used to consider the statistical uncertainty caused by a finite number of tests that is far from infinity.

177  $R_n$  in Eq. (3) can be calculated using the resistance prediction model  $g_R(\mathbf{x}_n)$  by inserting the nominal 178 input parameters  $\mathbf{x}_n$ . When nominal input parameters are not available, the characteristic values based 179 on the 5% fractile value can alternatively be used [29], or the nominal parameters can be inferred 180 from the product standard tolerance ranges.

181 The inverse analysis also uses the above-mentioned procedures, and the purpose is not to calibrate

the capacity factor for the given target reliability level, but to estimate the minimum required number of material tests to meet the target reliability level for a fixed/given capacity factor [24]. In particular, in the inverse analysis, the number of test samples (*n*) in Eq. (11) for material tests for the steel tensile strength is estimated. When the capacity factor ( $\phi$ ) is fixed or given, Eq. (3) takes the following form:

$$\phi R_n = R_d(n, \mathbf{x}) \tag{12}$$

where *n* is the only unknown term, and it can be numerically solved using optimisation algorithms
such as Active-Set Optimisation algorithm [30] and pattern search [31].

# **3. Forward and inverse reliability analyses for steel and composite columns**

The forward and inverse analysis methods presented in the previous section are conducted to estimate the capacity factors and the required minimum number of steel material tests for steel and composite columns, respectively. All the considered columns utilised helical seam pipe, or spirally welded steel tubes (SWTs). The following three types of sections are considered in these analyses: (i) steel hollow sections under pure compression, (ii) concrete-filled steel tubes (CFSTs) under pure compression, and (iii) CFSTs under eccentric loading (combined axial compression and uniaxial bending).

One concern of using SWTs in structural applications is whether anisotropy of the steel properties might exist, owing to the fact that the plate rolling direction and the longitudinal axis of the tube are no longer parallel. However, this concern has recently been addressed through two independent investigations on SWTs [32,33], where the results from tensile tests suggest that the material properties may be treated as isotropic. Due to this finding, in the remainder of this paper it is assumed that the yield strength of the steel in SWTs is independent of orientation.

## 201 **3.1 Parameter uncertainties**

202	All the parameters in the resistance prediction equation have aleatoric uncertainties, and in this study,
203	they are estimated from the manufacturing tolerances or the literature. The parametric uncertainties
204	are represented in terms of the bias and scatter, and the manufacturing tolerances for the tube diameter,
205	thickness, and the tensile strength of steel presented in Tables 1 and 2 are considered to estimate these.
206	We assume that these parameters are uniformly distributed within the tolerance ranges. It is assumed
207	that the mean of each parameter is the mid-point of the range, and the standard deviation of each
208	parameter that is uniformly distributed within the range of the minimum value of $r_{min}$ and the
209	maximum value of $r_{max}$ is calculated as $(r_{max} - r_{min})/\sqrt{12}$ . For all the other parameters that are not
210	specified in these manufacturing tolerances, the COV values have been taken from the literature such
211	that the COV of the compressive strength of concrete ( $f'_c$ ) is taken as 0.10 [34] and the COVs of all
212	linear dimensions unspecified in the manufacturing tolerances are taken as 0.01 [34].

Table 1. Manufacturing tolerances for hollow sections in API specification 5L [11]

Parameter	А	PI 5L
	$d_{\rm o} \leq 168.3$	-0.0075 <i>d</i> <sub>o</sub>
Outside dimension for CHS $(d_o)$ $(mm)$	$168.3 < d_0 \le 610$	-0.0075 <i>d</i> <sub>o</sub> (> -3.2)
	$d_{\rm o} > 610$	-0.005 <i>d</i> <sub>o</sub> (> -4.0)
	<i>t</i> ≤ 5.0	-0.5
Thickness for CHS (t) (mm)	$5.0 < t \le 15.0$	-0.1 t
	<i>t</i> > 15.0	-1.5

Table 2. Requirements for the results of tensile tests for PSL 2 pipes in API specification 5L [11]

Dine grade	Yield stren	ngth (MPa)
Pipe grade	Minimum value	Maximum value
X42 (290)	290	495
X46 (320)	320	525
X52 (360)	360	530
X56 (390)	390	545
X60 (415)	415	565
X65 (450)	450	600
X70 (485)	485	635

X80 (555)	555	705
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# **3.2 Steel hollow sections**

# *3.2.1 Experimental database*

219	Tests on 11 specimens have been collected for SWT columns to estimate the modelling uncertainty
220	statistically, and their input parameters together with their load bearing capacities are provided in
221	Table 3. In the table, $D$ = the diameter of the SWT, $t$ = the thickness of the tube, $f_y$ = the steel yield
222	strength, $L_e$ = the effective length of the column, and $P_{max}$ = the axial load bearing capacity of the
223	column. The experimental data were collected from the following four publications: Akiyama et al.
224	[35], Bao et al. [36], Gardner [37], and Aslani et al. [38]. These data are compared with resistance
225	model predictions, and the modelling error $\delta_i$ is statistically estimated using Eq. (6). It should be noted
226	that, in the tests by Aslani et al., the SWTs were manufactured with only single-sided MIG welding
227	on the outside of the pipe, as opposed to the requirement given in API 5L that the welding should be
228	both on the inside and outside. However, given the paucity of experimental data available on SWTs,
229	coupled with the fact that it is likely that any geometrical imperfections introduced from single-sided
230	welding would lead to more conservative resistances, the experimental data presented in Table 3 were
231	first assumed to belong to the same population within the reliability analyses. To investigate whether
232	this assumption was reasonable, the data was subsequently split into subsets corresponding to the
233	welding method and the reliability analyses repeated.

234	Table 3. Experime	ntal database for	spirally wel	ded steel tub	e columns u	nder axial co	mpression

Article	Specimen	<i>D</i> (mm)	<i>t</i> (mm)	$f_y$ (MPa)	L <sub>e</sub> (mm)	P <sub>max</sub> (kN)
Alainen et al. [25]	SA-7	400	7	409	1600	3360
Akiyama <i>et al.</i> [35]	SA-9	400	9	470	1600	5350
Dec. et al. [26]	c-st-1	323	6	344.35	970	1980
Bao <i>et al.</i> [36]	c-st-2	323	6	344.35	970	1990
Gardner [37]	9	168.8	2.64	298.12	1730	412.66

	10	169.3	2.64	317.74	1730	365.79
	H0SWT102-S	103.05	1.9	288	300	183.6
	H0SWT152-S	152.75	1.9	288	450	221.93
Aslani et al. [38]	H0SWT203-S	204.25	1.9	288	600	274.5
	H0SWT254-S	251.75	1.9	288	750	320.4
	H0SWT203-L	203	1.9	288	1400	296.8

## 236 *3.2.2 Resistance prediction models*

In NZS 3404.1 [19], AS 4100 [20], and AS/NZS 5100.6 [21], the nominal section capacity ( $N_s$ ) of a steel column under axial compression is predicted as follows:

$$N_s = \phi k_f A_n f_y \tag{13}$$

where  $\phi$  = the capacity factor for steel with a value of 0.9;  $k_f$  = the form factor defined as a ratio of the effective to the gross areas of the section;  $A_n$  = the nominal area of the steel sections; and  $f_y$  = the nominal yield strength of the structural steel.

242 The ultimate member capacity ( $N_c$ ) of a steel column under pure axial compression is predicted as 243 follows:

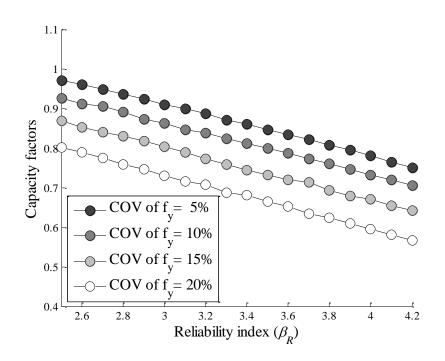
$$N_c = \alpha_c N_s \le N_s \tag{14}$$

where  $\alpha_c$  = the member slenderness reduction factor calculated according to NZS 3404.1, AS 4100, and AS/NZS 5100.6.

## 246 3.2.3 Analysis results

The calibrated capacity factors for steel in the circular hollow sections under axial loading are plotted in Figure 1. The forward analysis procedure introduced in the previous section is used. The capacity factors are calculated for the target reliability index  $\beta_R$  for resistance values, for the range of 2.5-4.2.

The analyses are repeated for different COV values of steel yield strength. As expected, the capacity 250 factor values decrease as the target reliability level increases and the COV of steel yield strength 251 increases. For the COV of steel yield strength of 7% [39] and  $\beta_R = 3.04$  (see Eq. (3)), the estimated 252 capacity factor has a value of 0.893 (see Table 4), which is close to the value of 0.90 provided in 253 NZS 3404.1, AS 4100, and AS/NZS 5100.6. This means that the capacity factor provided in the 254 current design standards just meets the desired target reliability level, but it does not have extra safety 255 or redundancy over the target reliability level. This is different from the design of other member types 256 or failure modes that usually have additional safety or conservatism by using capacity factors. This 257 258 is observed especially in the design of hollow section columns under axial compression, because the effect of the uncertainty in the section thickness directly and significantly affects the overall 259 parametric uncertainty of a member. This observation is consistent with the findings presented by 260 Kang et al. [24] for rectangular hollow sections under axial compression. 261



262

Figure 1. Calibrated capacity factor for steel in spirally welded steel tube columns under axial
 compression for various COV of the steel yield strength

To investigate whether it was reasonable to include the experimental data with single-sided welding, 266 the data was split into subsets corresponding to the welding method, and the reliability analyses were 267 repeated with a COV for steel yield strength of 7% and  $\beta_R = 3.04$  (see Table 4). To ensure a fair 268 comparison, the capacity factor  $\phi$  values are calculated by assuming that the fractile  $k_{d,m}$  evaluated for 269 the combined data (see Eq. (3)), remains constant in order not to introduce any additional statistical 270 uncertainty. As can be seen from Table 4, the values of bias and overall error are relatively insensitive 271 to the welding method. From this simple comparison, it appears that the initial assumption made in 272 Section 3.2.1 that the inclusion of the single-sided welding data would lead to more conservative 273 274 design resistances is reasonable, which is reflected in the lower capacity factor value in Table 4.

275 276

Table 4. Comparison of uncertainties and capacity factor for steel hollow sections under axial compression

	<i>n</i> (number	$\bar{b}$ (bias)	$V_{\delta}$ (Model	V <sub>Rt</sub>	$V_R$ (overall	$\phi$
	of tests)		error)	(Parametric	error)	
				error)		
All	11	0.993	0.064	0.083	0.105	0.893
Double-	6	0.984	0.068	0.083	0.108	0.906
sided						
welding						
Single-sided	5	1.004	0.064	0.084	0.105	0.861
welding						

277

Next, the inverse analysis according to the procedure provided in the previous section has been carried 278 out, and the minimum number of steel tensile tests required to meet the target value of  $\beta_R = 3.04$  have 279 been calculated. The analyses have been repeated for the COV of  $f_v$  values with the range of 2-12%. 280 281 The results are plotted in Figure 2, and it shows an increasing trend of the required number of material tests according to the increasing COV of  $f_y$ , as the uncertainty of the resistance prediction increases, 282 the statistical uncertainty due to the insufficient amount of material tests should be reduced to keep 283 the target reliability level. The proposed number of required steel material tests are also tabulated in 284 Table 5 according to this analysis for practical purposes. As can be seen, more than 30 tests are 285

required when the observed COV of the material tests is greater than 8%. This finding is because the resistance model has a very small conservatism and just meets the target reliability level even with infinite material tests.

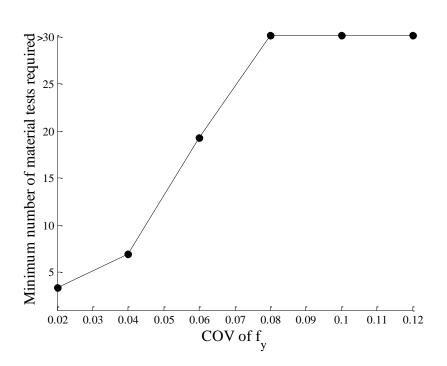


Figure 2. The minimum number of required steel material tests to meet the target reliability level for
 spirally welded steel hollow tubes under axial compression

292 293

294

289

Table 5. Recommended number of steel yield strength tests for spirally welded hollows sections under axial compression (based on  $f_{ym}/f_{yn} = 1.35$ )

COV of fy	2%	4%	6%	≥ 8%
# of tests	3	7	19	>30

295

# 296 **3.3 CFST sections under axial compression**

# 297 *3.3.1 Experimental database*

A total of 41 specimens have been collected for CFSTs with SWTs under axial compression. The input parameters and load bearing capacities are provided in Table 6, where  $f_{cm}$  is the mean measured concrete compressive strength. The experimental data were collected from the following six publications: Wang *et al.* [40], Aslani *et al.* [38], Akiyama *et al.* [35], Gardner [37], Gunawardena *et*  302 al. [41], and Gunawardena and Aslani [42]. However, in the tests by Aslani et al. together with 303 Gunawardena and Aslani, the SWTs were manufactured with only a single-sided MIG welding on the outside of the pipe (as opposed to the requirement given in API 5L that the welding should be 304 both on the inside and outside of the pipe). For the same reasons given in Section 3.2.1, the 305 experimental data presented in Table 6 were initially assumed to belong to the same population within 306 the reliability analyses. However, to investigate whether this assumption was reasonable, the data was 307 again split into subsets corresponding to the welding method and the reliability analyses subsequently 308 repeated. 309

For the cases when the mean measured cylinder compressive strength of concrete ( $f_{cm}$ ) values are not provided in the experimental data, but only the cube strength ( $f_{cu}$ ) are provided instead, the conversion table provided by Yu *et al.* [43] was used, which represents the approximate relationship between the cylinder strength ( $f_{cm}$ ) and the cube strength ( $f_{cu}$ ).

Table 6. Experimental database for CFSTs under axial compression

	-				f <sub>cm</sub>		<b>P</b> <sub>max</sub>
Article	Specimen	<b>D</b> (mm)	<i>t</i> (mm)	fy (MPa)	(MPa)	Le (mm)	(kN)
	CD4-1	425.8	5.2	259.8	42.51	1278	10523
	CD4-2	427.1	5.1	259.8	42.51	1278	10784
Ware at $a = \begin{bmatrix} 4 \\ 0 \end{bmatrix}$	CD6-1	628.5	6.9	276	42.51	1890	21207
Wang <i>et al.</i> [40]	CD6-2	628	7.1	276	42.51	1890	21582
	CD8-1	817.4	9	278.8	42.51	2460	36933
	CD8-2	820.8	9.3	278.8	42.51	2460	37221
	C-SWT102-S	103	1.9	288	38.78	300	609.81
	C-SWT152-S	152.25	1.9	288	38.78	450	1044.92
Aslani et al. [38]	C-SWT203-S	204	1.9	288	38.78	600	1788
	C-SWT254-S	252	1.9	288	38.78	750	2525
	C-SWT203-L	203.25	1.9	288	38.78	1400	1698
	SB20-7	400	7	409	21.7	1600	6740
Akiyama <i>et al.</i> [35]	SB20-9	400	9	470	25.6	1600	9510
	SB60-9	400	9	470	40.5	1600	12494
	la	168.8	2.64	298.12	17.95	305	1323.9
Gardner [37]	2a	168.8	2.64	298.12	34.13	305	1220.93
	3a	169.3	2.62	317.74	36.58	305	1304.28

	4a	169.3	2.62	317.74	33.54	305	1328.8
	5a	168.3	3.6	221.63	26.58	305	1559.26
	6a	168.3	3.6	221.63	32.75	305	1431.77
	6b	168.3	3.6	221.63	32.95	305	1461.19
	7a	168.8	5	260.86	32.95	305	1961.33
	7b	168.8	5	260.86	32.95	305	1966.23
	8a	168.8	5	260.86	27.46	305	2010.36
	8b	168.8	5	260.86	27.46	305	2010.36
	1	168.8	2.64	298.12	17.95	1830	823.76
	2	168.8	2.64	298.12	34.13	1830	916.92
	3	169.3	2.62	317.74	36.58	1830	757.07
	4	169.3	2.62	317.74	33.54	1830	690.39
	5	168.3	3.6	221.63	26.58	1830	947.32
	6	168.3	3.6	221.63	32.75	1830	1049.31
	7	168.8	5	260.86	32.95	1830	1132.67
	8	168.8	5	260.86	27.46	1830	1166.99
	LD1E0	102.67	1.83	234.9	24.20	1226	304
Gunawardena et al. [41]	LD2E0	152.74	1.76	234.9	23.05	1676	570
Gunawardena <i>et al</i> . [41]	LD3E0	203.04	1.93	234.9	24.34	2135	996
	LD4E0	229.81	1.96	234.9	24.22	2594	1228
	D1E0	102.9	1.83	234.9	29.9	614	363
C	D2E0	152.7	1.95	234.9	30.4	764	740
Gunawardena and Aslani [42]	D3E0	202.9	1.92	234.9	31.1	917	1249
	D4E0	230.0	1.94	234.9	29.3	1070	1495

## 316 *3.3.2 Resistance prediction models*

317 In AS/NZS 5100.6 [21] and AS/NZS 2327 [4], the design ultimate section capacity (N<sub>us</sub>) for a circular

318 CFST column under axial compression is given as:

$$N_{us} = \phi A_s \eta_2 f_y + \phi A_c f_c' \left( 1 + \frac{\eta_1 t f_y}{d_0 f_c'} \right)$$
(15)

where  $A_s$  and  $A_c$  = the areas of the steel and concrete sections, respectively;  $f_y$  = the nominal yield strength of the steel;  $f'_c$  = the characteristic compressive strength of the concrete;  $\phi$  and  $\phi_c$  = the capacity reduction factors for steel and concrete, respectively (with the existing target values given as 0.9 and 0.65); and  $\eta_1$  and  $\eta_2$  = the coefficients accounting for the confinement effect (where  $\eta_1$ represents the concrete strength increase, and  $\eta_2$  represents the steel strength reduction). 324 The design ultimate member capacity  $(N_{uc})$  of CFST is given as follows:

$$N_{uc} = \alpha_c N_{us} \le N_{us} \tag{16}$$

For a stub column defined by  $L_e/d_o$  or  $L_e/b \le 4$ , in which  $L_e$  = the effective length of a column and b = the section width of a rectangular tube,  $\alpha_c = 1$ .

327 3.3.3 Analysis results

328 The forward analysis results for CFSTs with SWTs under axial compression are presented in Figure 3. The analyses are repeated for the target reliability index  $\beta$  for resistance values, for the range of 329 2.5-4.2 and for different COV values of steel yield strength. The results are more sensitive to the 330 331 target reliability level as shown by the steeper slopes because there are two materials in CFSTs, and two types of capacity factors are included in the resistance model (i.e. one for steel and one for 332 concrete). When the concrete capacity factor is fixed, it is more difficult to change the reliability level 333 by changing the steel capacity factor alone compared to the case when there is only one capacity 334 factor in the whole design equation. Due to this issue, the increase of the COV of  $f_v$  does not 335 significantly reduce the capacity factor values because this uncertainty in  $f_y$  affects the steel 336 contribution only. The capacity factor value for steel with a COV of  $f_v = 7\%$  [39] is estimated as 1.194 337 when the capacity factor for concrete is fixed at 0.65 (see Table 7), which is greater than 0.90 provided 338 in AS/NZS 5100.6 [21] and AS/NZS 2327 [4]. This shows that the design model has extra safety 339 340 above the target reliability level, thereby providing a safe design.

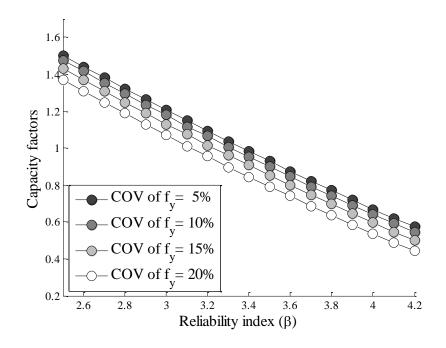




Figure 3. Calibrated capacity factor for steel in CFSTs with spirally welded steel tubes for various
 COV of the steel yield strength

To investigate whether it was reasonable to include the experimental data with single-sided welding, 345 the data was split into subsets corresponding to the welding method, and the reliability analyses were 346 repeated with a COV for steel yield strength of 7% and  $\beta_R = 3.04$  (see Table 7). Again, to ensure a 347 fair comparison, the capacity factor  $\phi$  values are calculated by assuming that the fractile  $k_{d,m}$  evaluated 348 for the combined data (see Eq. (3)), remains constant in order not to introduce any additional statistical 349 350 uncertainty. As can be seen from Table 7, the biases are overall similar for all data, double-sided welding, and single-sided welding. However, the modelling error of the single-sided welded 351 specimens is lower than the other specimens, which is considered to be caused by the random nature 352 353 of the data from the smaller sample size, together with better controlled experiments that were undertaken within the same laboratory. Notwithstanding this, the overestimation of the capacity factor 354 does not much affect the overall capacity factor and, the calculated capacity factor is already well 355 above the target value of 0.90 given in AS/NZS 5100.6 and AS/NZS 2327. It is concluded that the 356 addition of the single-sided welding data reduces the overall statistical error and the associated extra 357

			compres	551011		
	n (number	$\overline{b}$ (bias)	$V_{\delta}$ (M	odel $V_{Rt}$	$V_R$ (overall	$\phi$ (without
	of tests		error)	(Parametric	error)	extra
				error)		statistical
						error)
All	41	1.041	0.142	0.067	0.158	1.194
Double-	28	1.058	0.164	0.064	0.176	1.046
sided						
welding						
Single-sided	13	1.003	0.121	0.072	0.141	1.310
welding						

 Table 7. Comparison of uncertainties and capacity factor for CFST sections under axial compression

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Next, the inverse analyses have been carried out for CFST columns with SWTs. The results are shown 362 in Figure 4. As the design model for CFST columns has extra safety due to the conservatism that 363 exists within the equation and the capacity factors, the required material test numbers are quite low 364 to achieve the target reliability level for resistance with  $\beta_R = 3.04$ . The shallow slope of the graph 365 shows that the statistical uncertainty due to the number of material tests does not significantly affect 366 the overall reliability of the design, compared to other types of uncertainties such as parametric and 367 modelling uncertainties. The proposed required material test numbers are also tabulated in Table 8 368 for practical purposes. 369

370

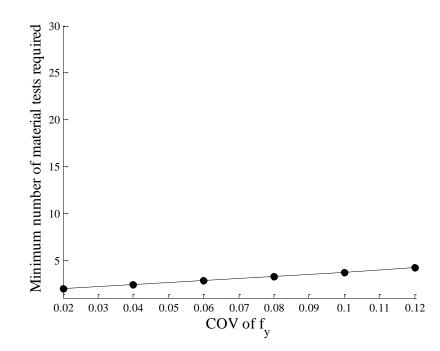


Figure 4. The minimum number of required steel material tests to meet the target reliability level for
 CFSTs under axial compression

Table 8. Recommended number of steel yield strength tests for CFST with spirally welded tubes under axial compression (based on  $f_{vm}/f_{vn} = 1.35$ )

COV of fy	2%	4%	6% 8%		10%	12%	
# of tests	2	2	3	3	4	4	

## **3.4 CFST sections under axial compression and uniaxial bending**

## *3.4.1 Experimental database*

A total of 16 specimens have been collected for CFSTs with SWTs under axial compression and uniaxial bending due to eccentric loading (see Table 9). The experimental data were collected from the following two publications: Gunawardena et al. [41] and Gunawardena and Aslani [42]. The uniaxial bending occurs due to the loads applied with the eccentricity of either at 0.15D or 0.40D. The SWTs for these tests were manufactured with only a single-sided MIG welding on the outside of the pipe (as opposed to the requirement given in API 5L that the welding should be both on the inside and outside of the pipe). However, from the comparisons of the uncertainties and capacity factors for the different welding methods presented in Section 3.2.3 and 3.3.2, coupled with the lack of 

388 experimental data available, it is assumed that the results are insensitive to the welding method.

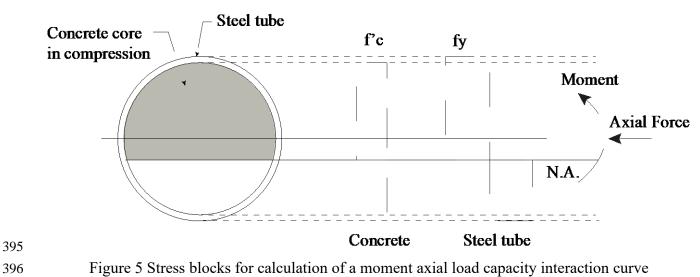
		D	t	$f_{y}$	<b>f</b> cm	Le	Mmax	<b>P</b> max	e
Article	Specimen	(mm)	(mm)	(MPa)	(MPa)	(mm)	(kNm)	(kN)	(mm)
	LD1E1	102.77	1.77	234.9	24.2	1226	5.8	254	0.15D
	LD1E2	102.72	1.83	234.9	24.18	1226	8.1	162	0.40D
	LD2E1	152.76	1.78	234.9	23.15	1676	13.7	421	0.15D
Gunawardena et	LD2E2	152.72	1.81	234.9	23.15	1676	20	269	0.40D
<i>al.</i> [41]	LD3E1	202.99	1.95	234.9	24.27	2135	25.4	601	0.15D
	LD3E2	202.15	1.93	234.9	24.34	2135	33.6	351	0.40D
	LD4E1	229.69	1.98	234.9	24.25	2594	33.9	723	0.15D
	LD4E2	230.14	1.95	234.9	24.25	2594	43.7	395	0.40D
	D1E1	102.9	1.88	234.9	29	614	4.8	298	0.15D
	D1E2	102.9	1.86	234.9	29.9	614	8.9	195	0.40D
	D2E1	152.7	2.1	234.9	30.4	764	15.1	557	0.15D
Gunawardena and	D2E2	153	1.99	234.9	30.4	764	21.7	327	0.40D
Aslani [42]	D3E1	202.8	1.86	234.9	31.1	917	27.1	778	0.15D
	D3E2	202.7	1.85	234.9	31.1	917	38.2	445	0.40D
	D4E1	229.5	1.86	234.9	29.7	1070	37.9	939	0.15D
	D4E2	229.7	1.93	234.9	29.3	1070	51.1	522	0.40D

Table 9. Experimental database for CFSTs under axial compression and uniaxial bending

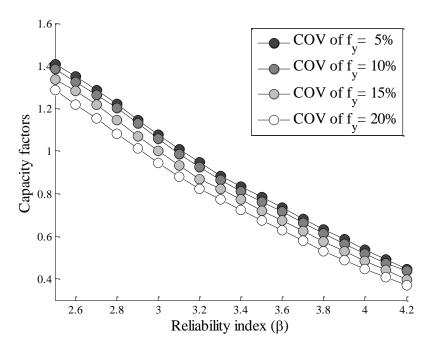
## 390 *3.4.2 Resistance prediction models*

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The axial and moment resistance of a CFST column under axial compression and uniaxial bending is derived from the moment axial load (M-N) interaction curves where  $M = N \cdot e_{eff}$ , considering the effective increase of eccentricity because of the 2<sup>nd</sup> order effects, and  $e_{eff}$  = effective eccentricity. The interaction curve is derived from the stress distributions along the section as shown in Figure 5.

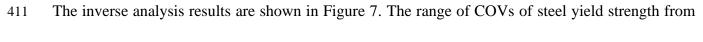


The forward analysis results for CFSTs with SWTs under axial compression and uniaxial bending are 398 presented in Figure 6. The forward analysis results show the reliability level similar to the CFSTs 399 400 under axial compression. The slope is steep, similar to the case of CFSTs under axial compression, which is due to the fixed concrete capacity factor that gives a constant amount of safety. A greater 401 change in a steel capacity factor is needed to change the overall reliability level. The capacity factor 402 value for steel for the COV of  $f_v = 7\%$  [39] is estimated as 1.081 when that for concrete is fixed at 403 404 0.65, which is greater than 0.90 provided in NZS 3404.1 [19], AS 4100 [20], AS/NZS 5100.6 [21] and AS/NZS 2327 [4]. This means that the design model provides extra safety against the target 405 406 reliability level.



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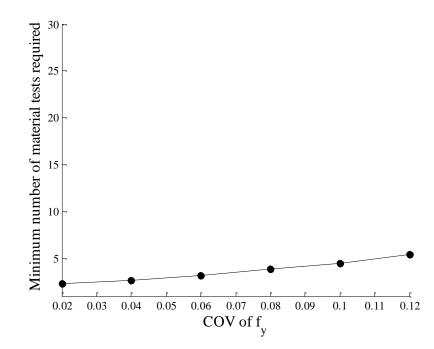
Figure 6. Calibrated capacity factor for steel in CFSTs with spirally welded steel tubes under axial
 compression and uniaxial bending for various COV of the steel yield strength



412 4% and 12 % have been considered. The trend is similar to CFST under axial compression, and it

requires a slightly greater number of material tests compared to that for CFST under axial compression. It is also shown that the effect of COV of  $f_y$  is similar to that of CFST under axial compression. The proposed numbers of required material tests are also tabulated in Table 10 for practical purposes.

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Figure 7. The minimum number of required steel material tests to meet the target reliability level for
 CFSTs under axial compression and uniaxial bending

422 Table 10. Recommended number of steel yield strength tests for CFST with spirally welded tubes 423 under eccentric loading (based on  $f_{ym}/f_{yn} = 1.35$ )

COV of fy	2%	4%	6%	8%	10%	12%
# of tests	2	3	3	4	4	5

424

# 425 **4. Conclusion**

This study carried out the capacity factor calibration for the design of steel and steel-concrete composite columns with helical seam pipe, which are also known as spirally welded steel tubes (SWTs). The analysis included the forward analysis to calibrate the capacity factor for steel, and the inverse reliability analysis to estimate the required number of material tests to meet the target reliability level for given capacity factors. This latter information will be of particular importance to
designers when the nominal yield strength of a product is unknown, as may occur when reclaimed
steel is reused. The following three types of members with circular cross-sections were considered:
(i) columns under axial compression, (ii) CFST columns under axial compression, and (iii) CFST
columns under combined axial compression and uniaxial bending due to eccentric loading.

For the SWT columns, 11 data were collected and used to estimate the modelling error. The design 435 equation gave the reliability just close to the target reliability index for resistance. This was because 436 of the effect of the uncertainty in thickness that directly affected the reliability and also the statistical 437 error due to the number of tests far from infinity. The inverse analysis results showed that the required 438 number of steel tensile tests quickly increased as the measured COV of  $f_y$  from the limited number of 439 material tests increased. For the CFST columns with SWTs under axial compression, 41 data were 440 collected and used to estimate the modelling error. The reliability achieved by the resistance model 441 was greater than the target reliability level showing the conservatism of the model. The inverse 442 analysis results showed that the required number of material tests did not quickly increase according 443 to the measured COV of  $f_v$  from the limited number of material tests, as the high reliability level was 444 already achieved by the design model itself. For the CFST columns with SWTs under axial 445 compression and uniaxial bending, 16 data were collected and used to estimate the modelling error. 446 The reliability achieved by the resistance model was similar to that of CFSTs under pure axial 447 compression and still higher than the target reliability level, showing the conservatism of the model. 448 The inverse analysis trend was similar to that of CFSTs under pure axial compression and it required 449 a slightly increased number of material tests to meet the target reliability level. 450

451 Some of the test specimens considered in this study consisted of SWTs manufactured with only 452 single-sided MIG welding on the outside of the pipe (as opposed to the requirement given in API 5L

that the welding should be both on the inside and outside of the pipe). From splitting the data into 453 subsets corresponding to the welding method, reliability analyses were repeated for SWTs and CFSTs 454 under axial compression. Whilst the test data is limited, the results suggest that the inclusion of test 455 specimens with single-sided welding leads to more conservative design resistances. However, the 456 calculated capacity factors generally still achieve similar or higher target values than provided by the 457 current Australian and New Zealand design standards. Notwithstanding this, the authors were unable 458 to find any test data for double-sided welded CFSTs subjected to combined axial compression and 459 uniaxial bending. As a consequence of this, further tests on members with different fabrication 460 461 methods are encouraged.

462

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