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# **Myopic Loss Aversion Revisited**\*

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#### **Abstract**

In this paper we reexamine several experimental papers on myopic loss aversion by analyzing individual rather than aggregate choice patterns. We find that the behavior of the majority of subjects is inconsistent with the hypothesis of myopic loss aversion.

**Key words**: myopic loss aversion, evaluation period, prospect theory

JEL Classification codes: D81, C91, D14

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## **Myopic Loss Aversion Revisited**

#### 1. Introduction

Myopic loss aversion (MLA) refers to a combination of greater sensitivity to losses than to gains (loss aversion) and a tendency to evaluate outcomes frequently (mental accounting) e.g. Thaler et al. (1997). Benartzi and Thaler (1995) proposed MLA as an explanation for the equity premium puzzle (Mehra and Prescott, 1985). However, Durand et al. (2004) showed that the analysis of Benartzi and Thaler (1995) is not robust. In a similar vein, Fielding and Stracca (2006) find that MLA can explain historical equity premium only if investors have highly short-sighted evaluation period.

Given these recent findings that macroeconomic simulations of equity premium puzzle do not appear to be consistent with MLA hypothesis, this papers takes a closer look at the experimental evidence in support of MLA. Thaler et al. (1997), Gneezy and Potters (1997), Gneezy et al. (2003), Langer and Weber (2005), Haigh and List (2005) and Bellemare et al. (2005) all find that, on average, subjects bet significantly higher amounts on a risky lottery when its performance is assessed over a relatively long time period. Thus, aggregate choice patterns observed in the above mentioned experiments appear to be in support of MLA. However, this paper shows that the majority of individual choice patterns in several of these experiments are actually inconsistent with MLA hypothesis.

The remainder of this paper is organized as follows. Section 2 presents the experimental design of Gneezy and Potters (1997), which was subsequently replicated (with additional treatments) by Langer and Weber (2005), Haigh and List (2005) and Bellemare et al. (2005). Section 3 reexamines the experimental results and shows that the majority of subjects exhibit behavior inconsistent with MLA. Section 4 concludes.

## 2. Experimental design

Langer and Weber (2005), Haigh and List (2005) and Bellemare et al. (2005) replicated the experiment of Gneezy and Potters (1997) and made several modifications to their design. Langer and Weber (2005) increased the number of rounds from 9 to 18 and used two other risky lotteries for which they did not find evidence of MLA. Haigh and List (2005) conducted a field experiment with professional traders from the Chicago Board of Trade. Bellemare et al. (2005) used an additional treatment identical to treatment L except that subjects observed the realization of the risky lottery in every period (betting behavior in this treatment was not significantly different from that in treatment H).

# 3. Reexamination of experimental results

Tables 1 and 2 show that the majority of subjects in the experiments of Gneezy and Potters (1997), Langer and Weber (2005) and Haigh and List (2005) exhibit the same individual choice patterns, both in treatment H and treatment L. In the majority of experimental rounds they invest an intermediate fraction of their initial endowment. Only a handful of subjects abstain from betting on the risky lottery and 12%-22% (15%-37%) of subjects consistently bet all their endowment on the risky lottery in treatment H (L).

	Number (percentage) of subjects					
Individual choice pattern in the majority of rounds <sup>1</sup>	Gneezy and Potters Weber (1997) (2005)		Haigh and List (2005), students	Haigh and List (2005), traders		
Invest 100% of endowment	7 (17.1 %)	2 (12.5%)	5 (15.7 %)	6 (22.2 %)		
Invest 1%-99% of endowment	27 (65.8 %)	13 (81.2%)	25 (78.1 %)	17 (63.0 %)		
Invest 0% of endowment	4 (9.8 %)	1 (6.3%)	1 (3.1 %)	2 (7.4 %)		
Other	3 (7.3 %)	0 (0.0%)	1 (3.1 %)	2 (7.4 %)		

Table 1 Individual choice patterns observed in Treatment H

	Number (percentage) of subjects				
Individual choice pattern in the majority of rounds	Gneezy and Potters Weber (1997) (2005)		Haigh and List (2005), students	Haigh and List (2005), traders	
Invest 100% of endowment	15 (35.7 %)	3 (15.0%)	6 (18.8 %)	10 (37.0 %)	
Invest 1%-99% of endowment	27 (64.3 %)	17 (85.0%)	26 (81.2 %)	17 (63.0 %)	
Invest 0% of endowment	0 (0.0 %)	0 (0.0%)	0 (0.0 %)	0 (0.0 %)	

Table 2 Individual choice patterns observed in Treatment L

Let us now focus on subjects who consistently bet an intermediate fraction of their endowment on the risky lottery. Table 3 shows that their intermediate bets are not significantly different across two treatments in all experiments with an exception of field experiment of Haigh and List (2005) with professional traders. We will now demonstrate that such behavior is inconsistent with the hypothesis of MLA.

The prediction of MLA is based on a deterministic decision theory proposed by Tversky and Kahneman (1992). According to their cumulative prospect theory, an individual derives utility from changes in wealth, which is captured by the value function  $v(x) = x^{\alpha}$  if  $x \ge 0$  and  $v(x) = -\lambda(-x)^{\beta}$  if x < 0. Coefficient  $\lambda > 0$  is the index of loss aversion (e.g. Köbberling and Wakker (2005)) and coefficients  $\alpha$  and  $\beta$ , that capture diminishing sensitivity to gains and losses, are estimated to be both equal to 0.88.

<sup>1</sup> Majority is defined as 5 rounds for experiments of Gneezy and Potters (1997) and Haigh and List (2005) and 10 rounds for the experiment of Langer and Weber (2005).

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Rounds	Average percentag (standard	Mann-Whitney					
	Treatment H	Treatment L	statistic ( <i>p</i> -value)				
<b>Gneezy and Potters (1997)</b>							
Rounds 1-3	43.71 (15.74)	50.00 (21.68)	-0.8118 (0.4169)				
Rounds 4-6	37.69 (19.74)	43.52 (16.34)	-1.2893 (0.1973)				
Rounds 7-9	45.40 (19.25)	56.24 (25.55)	-1.3390 (0.1806)				
Rounds 1-9	42.27 (15.58)	49.92 (16.54)	-1.7436 (0.0812)				
Langer and Weber (2005)							
Rounds 1-6	39.87 (24.29)	49.56 (16.49)	-1.5293 (0.1262)				
Rounds 7-12	39.87 (23.51)	52.94 (22.56)	-1.3404 (0.1801)				
<b>Rounds 13-18</b>	41.54 (26.38)	56.18 (25.71)	-1.3404 (0.1801)				
Rounds 1-18	40.43 (22.65)	52.89 (19.56)	-1.6744 (0.0940)				
Haigh and Lis	Haigh and List (2005), students						
Rounds 1-3	34.88 (20.49)	49.72 (21.59)	-2.5758 (0.0100)				
Rounds 4-6	46.13 (22.63)	52.28 (20.04)	-0.9641 (0.3350)				
Rounds 7-9	55.07 (25.02)	59.40 (23.61)	-0.6804 (0.4962)				
Rounds 1-9	45.36 (19.23)	53.80 (19.64)	-1.5922 (0.1113)				
Haigh and List (2005), professional traders							
Rounds 1-3	33.18 (23.69)	51.94 (21.90)	-2.3850 (0.0171)				
Rounds 4-6	27.98 (17.41)	60.69 (19.46)	-3.8802 (0.0001)				
Rounds 7-9	38.39 (26.28)	70.47 (22.31)	-3.2629 (0.0011)				
Rounds 1-9	33.18 (19.19)	61.03 (19.25)	-3.5493 (0.0004)				

Table 3 Average percentage of initial endowment invested in the risky lottery in treatments H and L by subjects who bet only intermediate fraction of their endowment in the majority of experimental rounds

An individual who invests nothing into the risky lottery obtains zero utility in both treatments. An individual betting amount x on the lottery in treatment H gets utility

(1) 
$$U_H(x) = (2.5x)^{\alpha} w_+(1/3) - \lambda x^{\beta} w_-(2/3)$$

where  $w_+(p) = p^{\gamma}/(p^{\gamma} + (1-p)^{\gamma})^{1/\gamma}$  and  $w_-(p) = p^{\delta}/(p^{\delta} + (1-p)^{\gamma})^{1/\delta}$  are the probability weighing functions for gains and losses respectively ( $p \in [0,1]$  and coefficients  $\gamma > 0$  and  $\delta > 0$  are estimated to be 0.61 and 0.69 correspondingly).

An individual betting amount x on the risky lottery in treatment L obtains utility

(2) 
$$U_L(x) = (0.5x)^{\alpha} w_+ (19/27) + (4^{\alpha} - 0.5^{\alpha}) x^{\alpha} w_+ (7/27) + (7.5^{\alpha} - 4^{\alpha}) x^{\alpha} w_+ (1/27) - \lambda (3x)^{\beta} w_- (8/27)$$

Let 
$$\underline{\lambda} = \frac{2.5^{\alpha} w_{+}(1/3)}{w_{-}(2/3)}$$
 and  $\overline{\lambda} = \frac{0.5^{\alpha} w_{+}(19/27) + (4^{\alpha} - 0.5^{\alpha})w_{+}(7/27) + (7.5^{\alpha} - 4^{\alpha})w_{+}(1/27)}{3^{\beta} w_{-}(8/27)}$ .

Notice that when  $\alpha=\beta$ , an individual bets nothing on the risky lottery in treatment H if her index of loss aversion  $\lambda$  is greater than  $\underline{\lambda}$  (in this case  $U_H(x)<0$ ). An individual bets all her initial endowment on the risky lottery if  $\lambda<\underline{\lambda}$  ( $U_H(x)>0$ ). Finally, an individual is exactly indifferent between betting and not betting (i.e. she can invest any fraction of her endowment in the risky lottery) if  $\lambda=\underline{\lambda}$  ( $U_H(x)=0$ ). Similar prediction holds for treatment L with the threshold for index of loss aversion being  $\overline{\lambda}$  instead of  $\underline{\lambda}$ .

For conventional parameterizations of cumulative prospect theory ratio  $\underline{\lambda}$  is smaller than ratio  $\overline{\lambda}$ . For example,  $\underline{\lambda} \approx 1.33$  and  $\overline{\lambda} \approx 1.66$  for parameters estimated by Tversky and Kahneman (1992).<sup>2</sup> Therefore, individual betting behavior in treatments H and L, which is consistent with the hypothesis of MLA, can be organized in the following Table 4 (depending on the unobservable index of loss aversion).

Index of loss aversion $\lambda$	$\lambda < \underline{\lambda}$	$\lambda = \underline{\lambda}$	$\underline{\lambda} < \lambda < \overline{\lambda}$	$\lambda = \overline{\lambda}$	$\lambda > \overline{\lambda}$
Betting on the risky lottery in treatment H	everything	anything	nothing	nothing	nothing
Betting on the risky lottery in treatment L	everything	everything	everything	anything	nothing

Table 4 Predicted behavior in treatments H and L according to MLA

In terms of the between-subject design of Gneezy and Potters (1997), Table 4 has the following testable implications:

A. Percentage of subjects, who bet all their endowment on the risky lottery, is higher in treatment L than in treatment H;

 $<sup>^2</sup>$  Gneezy and Potters (1997), Gneezy et a. (2003) and Haigh and List (2005) considered a simplified version of cumulative prospect theory without non-linear probability weighting ( $\gamma = \delta = 1$ ) and with a piecewise linear value function ( $\alpha = \beta = 1$ ). In this case  $\lambda = 1.25$  and  $\lambda \approx 1.56$ .

- B. Percentage of subjects, who abstain from betting, is higher in treatment H than in L;
- C. Percentage of subjects, who bet all their endowment in treatment L, is higher than the percentage of subjects, who bet an intermediate fraction of endowment in treatment H;
- D. Percentage of subjects, who bet nothing in treatment H, is higher than the percentage of subjects, who bet an intermediate fraction of their endowment in treatment L.

Tables 1 and 2 show that while implications A and B of MLA are confirmed for all experiments, implications C and D of MLA are clearly violated. In all experiments the majority of subjects decide to bet an intermediate fraction of their endowment on the risky lottery. The fraction of subjects who consistently bet an intermediate fraction of their endowment is nearly identical across two treatments (ranging between 65% and 85% across different experiments) and their intermediate bets are not significantly different across two treatments (except for the field experiment of Haigh and List (2005)).

MLA can explain this finding only if ratios  $\underline{\lambda}$  and  $\overline{\lambda}$  happen to be equal for the majority of subjects in both treatments. However, for the equality  $\underline{\lambda} = \overline{\lambda}$  to hold, we need to assume unconventional parameterizations of cumulative prospect theory (in particular,  $\delta < \gamma$ ), which contradicts to the existing experimental evidence (e.g. Tversky and Kahneman (1992), Abdellaoui (2000)). Moreover, if  $\underline{\lambda} = \overline{\lambda}$ , then MLA cannot explain implications A and B that apparently lead to statistically significant difference between aggregate choice patterns in treatments H and L.

#### 4. Conclusion

In this paper we reexamine the experimental evidence on risk taking and evaluation periods, provided by Gneezy and Potters (1997), Langer and Weber (2005) and Haigh and List (2005). A close look at the data suggests that behavioral patterns of the majority of subjects contradict to the MLA explanation. Subjects not only invest

intermediate fractions of their endowment, but also these intermediate bets do not appear to vary greatly across treatments with different length of evaluation period.

Our results suggest two important messages. First, experiments on risk taking and evaluation periods may reflect other phenomena than MLA. Some experimental evidence in support of the assumption of loss aversion of prospect theory has already been called into question by Plott and Zeiler (2005, 2006). They found that asymmetries in exchange behavior initially interpreted as evidence of endowment effect predicted by loss aversion are actually caused by subject misconceptions about the experimental procedure.

The second message is that the question of comparing expected utility theory and MLA approaches in the laboratory remains unanswered. While many attempts have been made to create an appropriate procedure, current experimental algorithms fail to discriminate between the two alternatives. It is left to further research to design an experiment, which would test expected utility theory versus MLA.

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